

Faculdade de Economia da Universidade de Coimbra

Grupo de Estudos Monetários e Financeiros (GEMF) Av. Dias da Silva, 165 – 3004-512 COIMBRA, PORTUGAL

gemf@fe.uc.pt
http://gemf.fe.uc.pt

JOSÉ A. SOARES DA FONSECA

The performance of the European Stock Markets: a time-varying Sharpe ratio approach

ESTUDOS DO GEMF

N.° 16 2009

PUBLICAÇÃO CO-FINANCIADA PELA FUNDAÇÃO PARA A CIÊNCIA E TECNOLOGIA

Impresso na Secção de Textos da FEUC COIMBRA 2009

The performance of the European Stock Markets: a time-varying Sharpe ratio approach

José A. Soares da Fonseca

Abstract

This article studies the performance of the national stock markets of sixteen European countries (Austria, Belgium, Denmark, Finland, France, Germany, Greece, Holland, Ireland, Italy, Norway, Portugal, Spain, Sweden Switzerland and United Kingdom), using daily data covering the period between 2nd January 2001 and 30th May 2009. Daily expected returns, and the conditional volatility of each index, were calculated using a model combining the market model and an implicit long-term relation between the index prices. Finally, time-varying (conditional) Sharpe ratios were calculated for each index. These were used as the basis for a statistical comparison of the performance of the stock indexes of this group of countries, throughout different sub periods corresponding to different conditions (of expansion and depression) in the stock markets.

Keywords: expected return, Sharpe ratio, market model, conditional volatility

JEL Classification: F36, G15

Introduction

This piece of research investigates the daily excess expected returns from sixteen European stock markets, and their conditional variance, in order to calculate time-varying Sharpe ratios, which are used to measure the performance of these stock markets between the beginning of 2001 and the middle of 2009. The use of these time-varying ratios allows a comparison between performance in different conditions (of growth and of contraction) for each market.

Simultaneously, these ratios are also used to evaluate the proximity of the performance between these countries under different market conditions. The stock markets under analysis, represented by their national stock indexes, are Austria, Belgium, Denmark, Finland, France, Germany, Greece, Holland, Ireland, Italy, Norway, Portugal, Spain, Sweden, Switzerland and United Kingdom.

In order to calculate daily time-varying Sharpe ratios for each market, we estimated the daily expected return and the conditional volatility of each market, using a model specified to include both a European market model, and an implicit long-term relation between the levels of the national and the European indexes. The estimations were carried out assuming the hypothesis that the volatility of the stock return follows a GARCH model from which the conditional volatility can be obtained. It is the joint predictability of the expected return and of the conditional volatility that allows the calculation of the time varying Sharpe ratios.

The inclusion of an implicit error correction model in the econometric procedure enables us to take into consideration methodology of financial integration analysis in which co-integration methods are used for the empirical analysis of stock market integration. On the other hand, the fact that the Sharpe ratios are calculated for a market portfolio, as is the case in this article, they can be defined as market prices of risk, in agreement with Leland (1999) and Adcock (2007). This also makes the methodology used in this article close to asset pricing models. In fact, in the approach to financial market integration based on the asset pricing models, which began with the seminal article of Solnik(1974), financial market integration is considered as being verified when the same asset pricing model can be applied to a group of domestic capital markets. The initial model of Solnik, which consisted of a world capital asset pricing model containing a world market price of risk, was later taken further by other authors, such as Stehle (1977) Jorion and Schwartz (1986) to include both a domestic and a world market price of risk. The hypothesis of market efficiency contained in capital asset pricing models has caused problems in the empirical analysis based on these models, because it is often contradicted by empirical results. This is one of the reasons why, in some more recent research, co-integration models have become popular in the empirical analysis of financial market integration. Cointegration provides a tool for measuring the interdependence between a domestic stock market and an international stock market both in the long- and short-terms. Additionally, co-integration models also take into account the influence exerted by lagged changes of the variables over their current changes, which is observed in the cases in which market efficiency is absent. First studies on the subject of European stock market integration using the co-integration approach were published early in this decade. Rangvid (2001) and Miloudi (2003) used cointegration methods as a tool for evaluating the integration of the European stock markets in the years before the launch of the single currency. Other studies, such as those of Kasa (1992), Arshanapalli and Doukas (1993), and Richards (1995) also applied co-integration to evaluate the integration of non-European stock markets.

The econometric method used and the theoretical background for the calculation of the time varying Sharpe ratios

In this research each national stock market is represented by its national MSCI (Morgan Stanley Capital International) Index, expressed in euros, and using daily data which covers the period between 1st January 2001 and 31st May 2009, and comprises 2195 observations of each national index. The European Index (MSCI) and the European Overnight Interest Average (EONIA) are the two other variables used in this research, also using daily data and covering the same period as the others. Prior to econometrical testing, each index series was transformed giving the base 100 on 2nd January 2001 for all the series.

The logs of these new series were consequently calculated and used in the estimations.

The model on which the estimation of the expected returns for each of the national index is based combines a European market model, and the long-term relation between the national index and the European index. The representation of the European market model is given by:

$$R_{i,t} = \alpha_i + \beta_i R_{F,t} + \varepsilon_{i,t} \tag{1}$$

where $R_{i,t}$ and $R_{E,t}$ are the return of the national portfolio and the return of the European portfolio over period t respectively, and $\varepsilon_{i,t}$ is the error term, which has, by hypothesis, a zero mean. Taking the operators of mathematical expectations, the representation of the market model becomes:

$$E_{t}(R_{i}) = \alpha_{i} + \beta_{i}E_{t}(R_{E})$$
(2)

where $E_t(R_t)$ is the expected return of the domestic portfolio (index) over period t, and $E_t(R_E)$ is the expected return of the European portfolio (index) also over period t.

The inclusion of the long-term relation between the national index and the European index is based on the error correction model of Engle-Granger (1987). Our tests were conducted using the logs of the index prices, which, from now on, will be represented in this paper by $p_i = \log(P_i)$. Thus, the error correction model takes the following form:

$$\Delta p_{i,t} = a_{1t} + a_{i,e} \left(p_{i,t-1} - \varphi_0 - \varphi_1 p_{E,t-1} \right) + \sum_{j=1}^{L} a_{11,j} \Delta p_{i,t-j} + \sum_{j=1}^{L} a_{12,j} \Delta p_{E,t-j} + \varepsilon_{i,t}$$
 (3).

which means that the current change in the price log of the i index at period t, $\Delta p_{i,i}$, is explained by the lagged deviation of its value relative to

the long-term relation with the log of European index, and by L lagged changes of the price logs of both of the domestic and the European indexes. As the changes in the price logs are the returns of the portfolios, the error correction model can take the following form:

$$R_{i,t} = a_{1t} + a_{i,e} \left(p_{i,t-1} - \phi_0 - \phi_1 p_{E,t} \right) + \sum_{i=1}^{L} a_{11,j} R_{i,t-j} + \sum_{i=1}^{L} a_{12,j} R_{E,t-j} + \varepsilon_{it}$$
 (4)

In the empirical analysis conducted in this article the hypothesis that the returns of a national index are determined by twice the influence of the market model, and of the error correction model, is tested. The combination of both influences are given by the following:

$$R_{i,t} = \varpi_1 \left[\alpha_i + \beta_i R_{E,t} \right]$$

$$+ \varpi_2 \left[a_1 + a_{i,e} \left(p_{i,t-1} - \varphi_0 - \varphi_1 p_{E,t} \right) + \sum_{j=1}^{L} a_{11,j} R_{i,t-j} + \sum_{j=1}^{L} a_{12,j} R_{E,t-j} \right] + \varepsilon_{it}$$
(5)

where ω_1 and ω_2 are the weights, respectively of the market model and of the error correction model, in the explanation of the daily return of the national index. The following equation was assigned to this model for econometrical estimation:

$$R_{i,t} = \alpha_i^* + \beta_i^* R_{E,t} + \varphi_1^* p_{i,t-1} + \varphi_2^* p_{E,t-1} + \sum_{i=1}^L a_{11,j}^* R_{i,t-j} + \sum_{i=1}^L a_{21,j}^* R_{E,t-j} + \varepsilon_{it}$$
 (6)

As Adcock (2007) notes, it is common practice to embed the beta (market) model in models with auto-regressive and/or moving average terms, which also take in consideration the hypothesis of ARCH/GARCH effects. That is the case of the model tested in the present piece of research. The main advantage of this econometrical procedure is that it makes evident, simultaneously, and through the

estimates of the coefficients, the importance of the European market model in the explanation of the daily returns of each national index, and the influence exerted by the prices or the lagged returns. The hypothesis that the conditional variance follows a GARCH model has also been considered in the tests. Thus, the estimation was made via a maximum likelihood procedure. The results of the tests confirmed that it is adequate to represent the conditional variance for all the national indexes under analysis using the GARCH(1,1)model:

$$\sigma_t^2 = \alpha_{\varepsilon} + \beta_{1,\varepsilon} \varepsilon_{t-1}^2 + \beta_{2,\varepsilon} \sigma_{t-1}^2 \tag{7}$$

(where σ_{t^2} is the conditional variance at time t, and ε_{t-1} is the error term squared).

After the estimation, the normalized residuals (i.e. the residuals divided by the square root of the conditional variance) were tested for autocorrelation, using a Ljung-Box test, and for ARCH, using an F test on the coefficients of an autoregressive model of the squared normalized residuals:

$$\varepsilon_t^2 = a + \sum_{j=1}^k b_j \varepsilon_{t-j}^2 \tag{8}$$

Both the Ljung-Box test and the ARCH test were carried out for a maximum of 24 lags, with a span of 4 lags. The results of these two tests determined the choice of the number of lags in the mean equation, and also the type and the order of the GARCH model of the conditional variance. According to the results of these tests, as will be discussed in more detail later, one lag (L=1) in the mean equation has been shown to in almost all the eliminate residual adequate cases to autocorrelation. The only exception was the case of Sweden, in which it was necessary to include two lags of the dependent variable in the mean equation in order to eliminate the autocorrelation of the residuals.

One of the primary uses of the expected returns, $E(R_i)$ and of the risk, σ_t , is to calculate the Sharpe ratio:

$$S_i = \frac{E(R_i) - r_f}{\sigma} \tag{9}$$

where r_f is the return of the risk free asset. The calculation of this ratio allows a comparison between the performances of the stock market of country i and the stock markets of other countries. Leland (1999) and Adcock (2007) defined this Sharpe ratio, when related to a stock market, as the *market price of risk*. Both Leland and Adcock based their analysis on the non conditional CAPM, which implies that the market price of risk is constant during the period covering the data used to calculate the expected return and the risk.

As the empirical model estimated in the present piece of research produces daily time varying expected returns $E_t(R_i)$, and a time-varying measure of risk, the conditional volatility $\sigma_{i,t}$, a daily time varying Sharpe ratio, as shown by the following expression:

$$S_{i,t} = \frac{E_t(R_i) - r_{f,t}}{\sigma_{i,t}} \tag{10}$$

can also be calculated for each national index, (the risk-free interest rate used in the calculation is the European Overnight Interest Average).

The use of a stochastic discount factor as a tool for asset pricing forms the theoretical basis for the economic interpretation of the timevarying Sharpe ratio. In a non-arbitrage economy with complete markets all the assets can be priced using the stochastic discount factor (or pricing kernel) of the Harrison and Kreps (1979) type , M_{t+1} , which satisfies the following condition for any asset, or portfolio i:

$$E_{t}(M_{t+1}R_{i,t+1}) = 1 (11)$$

where $R_{i,t+1}=log(P_{i,t+1}/P_{i,t})$

In agreement with the non-arbitrage condition, equation (11) can also be applied to the risk-free asset, which can, thus, be represented by the inverse of the expectation of the pricing kernel:

$$r_{f,t} = E_t \left(M_{t+1} \right)^{-1} \tag{12}$$

Developing Equation (11) in accordance with the rules of the expectation of the product of two random variables, and replacing $E_t(M_{t+1})^{-1}$ by $r_{f,t}$, it can be concluded that the excess expected return of the portfolio i is proportional to its conditional covariance with the pricing kernel, i.e:

$$E_{t}(R_{i,t+1}) - r_{f,t} = -r_{f,t}Cov_{t}(M_{t+1}, R_{i,t+1})$$
 (13)

where Cov_t is the conditional covariance. Dividing equation (13) by the conditional standard deviation of the portfolio i, $\sigma_{i,t}$, it is possible to conclude that the conditional Sharpe ratio of the portfolio i is proportional to the conditional correlation between the return of the portfolio and the pricing kernel:

$$S_{i,t} = -r_{f,t} \sigma_{M,t} Corr_t (M_{t+1}, R_{i,t+1})$$
 (14)

where $\sigma_{M,t}$ is the conditional standard deviation of the pricing kernel, and Corr_t is the conditional correlation between it and portolio i. As Whitelaw (1994, 1997) underlines, we can intuitively conclude that a substantial part of the variation of the conditional Sharpe ratio is attributable to variation in this conditional correlation. On the same lines as Whitelaw, goes the empirical evidence of Ayadi and Krysanovsky (2008), that the use of pricing kernel methodology can easily encompass time-varying measures of performance. Both the

postulate of Whitelaw, and the empirical evidence of Ayadi and Krysanovsky show the importance of calculating time-varying Sharpe ratios as they provide an indirect way of obtaining information regarding the conditional correlation between the return of a market portfolio and the stochastic discount function (or, in a similar way, on the conditional correlation between the return of a market portfolio and the variables affecting the stochastic discount function).

The final objective of this article is to evaluate the co-movement of the conditional Sharpe ratios of this group of national indexes. The use of historical correlation is a possible tool for this objective. However, it is not suitable for taking into account the possibility that the correlations change over time. Thus, it was used the *cross-sectional dispersion measure*, proposed by Solnik and Roullet (2000), initially to be applied to stock returns, which varies inversely with instantaneous average correlation, and so provides information regarding dynamic correlation. This measure, applied in this paper, is represented by the variance across the national index Sharpe ratios, and was calculated daily. Its representation, referred to each period t:

$$CSDM_{t} = \sum_{i=1}^{16} \left(S_{i,t} - \overline{S}_{t} \right)^{2}$$
 (15)

where \overline{S}_t is the average Sharpe ratio over period t.

The statistical analysis of the series of the CSDM, through different subsamples of the period under analysis, gives information regarding the inter temporal evolution of the proximity of the performance of the indexes under analysis. We can take the proximity of the Sharpe ratios as an indicator of the degree of integration of the financial markets. Thus, conducting statistical tests on the CSDM over different subsamples, we arrive at conclusions regarding the evolution of the integration within the group of domestic financial markets. These tests were conducted on the series of the CSDM referring to these 16 countries, and, separately, the same tests were applied to the eleven euro area countries. Since the subsamples considered in these tests

correspond to different phases of the stock market, it was possible to arrive at a comparative analysis of the integration of these markets in phases of both financial market expansion and contraction.

The estimation of the expected returns, Sharpe ratios and analysis of its evolution

The results of the estimation of the combined market model-error correction model, and the GARCH, for each of the stock indexes are shown in Tables I.1 to I.16. Each of these refers to one of the national indexes under study. Each table is composed of three separate parts. In the first part, a), the results of the estimation of the mean equation and the GARCH model are represented. These include, for each coefficient, the estimate, the standard error, the T statistic and the significance level. In the second part, b), results (the Chi-squared test statistic and the significance level) of the Ljung-Box tests on the autocorrelation of the residuals are shown. These refer to a maximum of 24 lags with a span of 4 lags. In the third part the tests on the residuals heteroskedasticity (ARCH), which consist on the F test statistic and (its level of significance) calculated through the estimation of autoregressive models of the squared residuals with a maximum of 24 lags and a span of 4 lags are given.

The results presented in these tables show that, in the explanation of the daily returns of major part of the national indexes, the market model dominates the influence exerted by the national and the European index values, since, for all the countries, the coefficient of the return of the European index is significantly different from zero. On the other hand, in the majority of the cases, the coefficients of the national and the European index values are not significantly different from zero. The exceptions to this rule are the cases of Finland, France, Portugal and Switzerland. In these cases the statistics of the coefficients of the national, and the European indexes, lead to the rejection of the

null hypothesis that they are not significantly different from zero. Since the coefficients of the index values contain information regarding the long-term relation between each national index and the European index, it can be taken that, in the case of these four countries, the return of their national stock indexes is explained both by a European market model and by the implicit long-term relation between the national index and the European index. The German case is peculiar because the coefficient of the European index level is significantly different from zero, while the opposite situation is observed with the coefficient of the domestic index.

According to the results of the Ljung-Box test, shown in part b) of Tables I.1 to I.16, and also according to the results of the ARCH test, in part c) of those tables, there is no autocorrelation nor ARCH effects observed in the residuals of any of the regressions.

As mentioned above, the second part of the tests conducted for this article involved the calculation of daily Sharpe ratios for each national index, and their statistical analysis, both over the total period of analysis, and over different subsamples. The total period, between 1st 31st May 2009, was broken down into four January 2001 and subsamples: 1) between 1st January 2001 and 31st December 2002, 2) between 1st January 2003 and 31st December 2004, 3) between 1st January 2005 and 31st December 2006, and 4) between 1st January 2007 and 31st May 2009. During the first and fourth subsamples phases of market contraction were predominant, while during the second and the third periods the financial markets predominantly went through phases of growth (This is illustrated in Figure 1, where the series of the European index is given). The main statistics on the timevarying Sharpe ratio of each country, relative to the entire period and to the four subsamples are presented at the Table II. In general, the average of the time-varying Sharpe ratios is positive in the subsamples during which the stock markets predominantly experienced phases of growth. On the contrary, in the subsamples during which the decrease in prices was dominant, the average of the conditional Sharpe ratio is

negative. The Sharpe ratio is negative when the index expected return is less than the risk-free interest rate. This situation is not necessarily precluded by the equilibrium situation in the stock market, if, as Boudoukh, Richardson and Whitelaw (1997) found, there is a nonlinear relation between the equity risk premium and the slope of the term structure of interest rates.

These statistics (mean, standard error and level of significance) are complemented by a test for equality across the subsamples. The results of this test represented by the Chi-squared statistics and the respective level of significance, presented together with the other results of each national index, confirm that the behaviour of the Sharpe ratios was not equal across subsamples.

The ex-post Sharpe ratio,:

$$S_{i}^{EP} = \frac{\mu}{\sqrt{\sum_{t=1}^{T} \left[\left(R_{i,t} - r_{f,t} \right) - \mu \right]^{2}}}$$

$$T - 1$$
(14)

where $\mu = \frac{\sum\limits_{t=1}^T \left(R_{i,t} - r_{f,t}\right)}{T}$ and T is the number of observations, was calculated for the whole sample, and for the subsamples. The ex-post Sharpe ratio has, in every case, the same sign as the average conditional Sharpe ratio, as it is also shown in Table II.

The statistics regarding the series of the cross section dispersion measure (CSDM) of the conditional Sharpe ratios, between the 16 national stock indexes under analysis, are given in Table III. These statistics were calculated for the entire period as well as for the four subsamples referred to previously. These statistics (mean, standard error and level of significance) were also complemented with a test for equality across the subsamples. The results of this test, represented by the Chi-squared statistics and the respective level of significance, are

also given in Table III. The average CSDM shows the lowest average value in the subsample relating to 2003-2004, which was dominated by periods of growth in the stock markets, and the highest average value in the last subsample, relating to 2007-2009, which mostly corresponds to the period following the 2007 financial crises. Figure II shows the CSDM series and illustrates these conclusions. The fact that an increase in the CSDM was particularly notable during the period following the 2007 crises suggests that there was an intensive increase in domestic bias after the crises, which is, quite probably, one of the main causes of the reduced degree of integration.

The CSDM was also calculated for the Sharpe ratios of the eleven EMU member countries (Austria, Belgium, Finland, France, Germany, Greece, Holland, Ireland, Italy, Portugal and Spain) and the statistical tests, which are given in Table IV and illustrated graphically in Figure III, lead to conclusions similar to those obtained for the complete group of sixteen countries. The average CSDM, observed over the last subsample was remarkably higher than those observed over the other subsamples. This result can be interpreted as meaning that, even within the stock markets of the EMU members, the 2007 crises caused a reduction in their degree of integration.

Conclusions

The empirical analysis conducted in this article shows that time-varying Sharpe ratios are an adequate tool for a comparative analysis of the performance of different stock markets, and also that they help us to have a perspective on the dynamics of their integration. To calculate the time-varying Sharpe ratios for sixteen European stock indexes, the conditional mean and the conditional volatility of the indexes were estimated by a model whose specification combined the market model and the influence of the long-term relation between each national index and the European index. The results of these estimations showed that the market model component is dominant, obscuring the

influence of the implicit long-term relation between the national and the European index in almost all cases. The exceptions to this rule were the cases of Finland, France, Portugal and Switzerland, in which, there was evidence of the explanatory power of the index levels.

The statistical analysis of the conditional Sharpe ratios showed that they present, on average, clear differences between the growth phases (during which higher performance was observed) and the depression phases of the stock market (during which lower performance dominated).

Finally, the calculation of a cross dispersion measure, both across the group of sixteen countries and across the EMU members only, showed that the dispersion of the performance experienced a much more significant increase over the period following the 2007 crisis than that observed in the years preceding it. This result can be interpreted as evidence that the 2007 crisis caused a negative break in the process of integration between the markets under analysis.

References

Adcock. C. (2007), "Measuring portfolio performance using a modified measure of risk", *Journal of Asset Management*, Vol. 7, 388-403.

Ayadi, M. and Krysanovsky, L. (2008), "Portfolio performance sensivity for various asset, C pricing kernels", Computers & Operations Research, 35, pp 171-185.

Arshanapalli, B. and Doukas, J. (1993), "International stock market linkages: Evidence from the pre- and post-October 1987 period", *Journal of Banking and Finance*, 17, pp.193-208.

Boudoukh, J., Richardson, M. And Whitelaw (1997), "Nonlinearities in the Relation Between the Equity Risk Premium and the Term Structure", *Management Science*, 43, pp.371-385.

Engle, R. and Granger, C. (1987), "Cointegration and Error-Correction: Representation, Estimation and Testing", *Econometrica*, N° 55, pp. 251-276.

Harrison, M. and Kreps, D. (1979), "Martingales and Arbitrage in Multiperiod Secutity Markets", *Journal of Economic Theory*, 20, 381-408.

Jorion, P. and Schwartz, E. (1986), "Integration versus Segmentation in the Canadian Stock Market", *The Journal of Finance*, Vol. XLI, N°3, pp.603-616.

Kasa, K. (1992) "Common stochastic trends in international stocks markets", Journal of Monetary Economics, 29, pp 95-124.

Leland, H. (1999) "Beyond Mean-Variance: Performance Measurement in a Nonsymmetrical World", *Financial Analysts Journal*, Jan-Feb.pp. 27-36.

Miloudi, A. (2003), "Interdépendances entre Places Financières Européennes: une Analyse en terme de Cointégration et de Causalité", document de recherche, ATER en Finance, Université de Rennes.

Rangvid, J. (2001), "Increasing convergence among European stock markets? A recursive common stochastic trends analysis", *Economics Letters*, 71, pp.383-389.

Richards, A. (1995), "Co movements in national stock markets returns: Evidence of predictability, but not co integration", *Journal of Monetary Economics*, 36, pp. 631-654.

Stehle, R. (1977), "An Empirical Test of the Alternative Hypothesis of National and International Pricing of Risky Assets", *The Journal of Finance*, Vol. XXXII, N°2, pp.493-502

Solnik, B. (1974), "An Equilibrium Model of International Capital Market", *Journal of Economic Theory*, N° 8, pp. 500-524.

Solnik, B. and Roulet, J. (2000), "Dispersion as a Cross-Sectional Correlation", *Financial Analysts Journal*, January-February, pp. 54-61

Whitelaw, R. (1994), "Time Variations and Covariations in the Expectation and Volatility of Stock Market Returns", *Journal of Finance*, 49, pp. 515-541.

Whitelaw, R. (1997), "Time-Varying Sharpe Ratios Market Timing", Working Paper University of New York.

Table I.1: Estimation of the conditional mean return and conditional volatility

Austria

a) Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
$lpha_i^*$	0.00572730	0.00381320	1.50195000	0.13310978
$\boldsymbol{\beta}_i^*$	0.53160000	0.02330000	22.85267000	0.00000000
ϕ_1^*	-0.00062654	0.00057490	-1.08983000	0.27578804
ϕ_2^*	-0.00039888	0.00125950	-0.31669000	0.75147554
a_{11}^*	-0.01400000	0.02150000	-0.64892000	0.51638901
a_{21}^*	0.08460000	0.02310000	3.65682000	0.00025536
		GARCH(1,1)		
$lpha_arepsilon$	0.00000179	0.00000048	3.71118000	0.00020630
$oldsymbol{eta}_{\!\scriptscriptstyle 1,arepsilon}$	0.0880000	0.01250000	7.01157000	0.00000000
$oldsymbol{eta}_{2,arepsilon}$	0.89690000	0.01440000	62.44657000	0.00000000

b)The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	0.7169	0.94923
LB(8)	5.1135	0.74537
LB(12)	9.6957	0.64264
LB(16)	15.6285	0.47917
LB(20)	17.8521	0.59714
LB(24)	18.5478	0.77565

	Test Statistic	Significance Level
ARCH(4)	0.17593	0.95084
ARCH(8)	0.64351	0.74153
ARCH(12)	0.81718	0.63292
ARCH(16)	0.98474	0.47057
ARCH(20)	0.89672	0.59170
ARCH(24)	0.78294	0.76233

Table I.2: Estimation of the conditional mean return and conditional volatility

Belgiuma) Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
$lpha_i^*$	0.00695860	0.00273020	2.54876000	0.01081076
$oldsymbol{eta}_i^*$	0.81080000	0.01400000	57.71465000	0.00000000
ϕ_1^*	-0.00105720	0.00128820	-0.82062000	0.41186309
ϕ_2^*	-0.00043997	0.00174010	-0.25284000	0.80038934
a_{11}^*	0.01140000	0.02600000	0.44087000	0.65930594
a_{21}^*	0.03110000	0.02450000	1.27273000	0.20311267
		GARCH(1.1)		
$lpha_{arepsilon}$	0.00000063	0.0000013	4.65277000	0.00000328
$oldsymbol{eta}_{1,arepsilon}$	0.09190000	0.01240000	7.40991000	0.00000000
$oldsymbol{eta}_{2,arepsilon}$	0.90240000	0.01160000	77.68092000	0.00000000

b)The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	3.6698	0.452536
LB(8)	6.0131	0.645766
LB(12)	7.7285	0.805969
LB(16)	8.9035	0.917332
LB(20)	13.2755	0.865251
LB(24)	17.6332	0.820431

c)F-Test of no ARCH vs. ARCH in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
ARCH(4)	0.88339	0.472981
ARCH(8)	0.72242	0.67190
ARCH(12)	0.67190	0.81231
ARCH(16)	0.53542	0.92979
ARCH(20)	0.64304	0.88263
ARCH(24)	0.72967	0.82518

Table I.3: Estimation of the conditional mean return and conditional volatility

Denmark

a) Coefficients of the conditional mean and conditional volatility

	•				
Coeff	Estimate	Std Error	T-Stat	Signif	
$lpha_i^*$	-0.00172060	0.00369990	-0.46505000	0.64189909	
$oldsymbol{eta}_i^*$	0.66620000	0.01530000	43.43320000	0.00000000	
ϕ_1^*	-0.00085770	0.00111420	-0.76978000	0.44143161	
ϕ_2^*	0.00140400	0.00174710	0.80365000	0.42159961	
a_{11}^*	-0.03980000	0.02130000	-1.86475000	0.06221661	
a_{21}^*	0.11330000	0.02070000	5.48065000	0.00000004	
	GARCH(1.1)				
$lpha_arepsilon$	0.00000078	0.00000037	2.10445000	0.03533921	
$oldsymbol{eta}_{1,arepsilon}$	0.05170000	0.01250000	4.13850000	0.00003496	
$oldsymbol{eta}_{2,arepsilon}$	0.93930000	0.01520000	61.73197000	0.00000000	

b)The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	1.7903	0.774259
LB(8)	16.8255	0.031979
LB(12)	19.0128	0.088219
LB(16)	23.9165	0.091344
LB(20)	28.0496	0.10823
LB(24)	31.1722	0.148879

	Test Statistic	Significance Level
ARCH(4)	0.44149	0.77868
ARCH(8)	2.11425	0.03147
ARCH(12)	1.66286	0.06883
ARCH(16)	1.45648	0.10711
ARCH(20)	1.37645	0.12258
ARCH(24)	1.22137	0.21037

Table I.4: Estimation of the conditional mean return and conditional volatility

Finland

a)Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
α_i^*	-0.00933624	0.00571014	-1.63503000	0.10204312
$oldsymbol{eta}_i^*$	1.13403301	0.02143894	52.89595000	0.00000000
ϕ_1^*	-0.00769559	0.00242934	-3.16777000	0.00153611
ϕ_2^*	0.00887745	0.00285477	3.10969000	0.00187282
a_{11}^*	0.03150981	0.02347229	1.34243000	0.17945794
a_{21}^*	-0.00008580	0.03387621	-0.00253000	0.99797906
		GARCH(1.1)		
$lpha_{arepsilon}$	0.00000020	0.00000007	3.05245000	0.00226983
$oldsymbol{eta}_{1,arepsilon}$	0.01104221	0.00153801	7.17956000	0.00000000
$oldsymbol{eta}_{2,arepsilon}$	0.98740983	0.00146830	672.48744000	0.00000000

b)The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	7.4716	0.112969
LB(8)	12.1906	0.142899
LB(12)	19.9516	0.068007
LB(16)	24.0749	0.087881
LB(20)	29.6503	0.075723
LB(24)	31.5268	0.139167

	Test Statistic	Significance Level
ARCH(4)	1.71003	0.14500
ARCH(8)	1.46346	0.16546
ARCH(12)	1.64414	0.07326
ARCH(16)	1.55108	0.07425
ARCH(20)	1.58776	0.04711
ARCH(24)	1.34646	0.12119

Table I.5: Estimation of the conditional mean return and conditional volatility

France

a)Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
α_i^*	0.00100098	0.00136941	0.73096000	0.46480359
$oldsymbol{eta}_i^*$	1.06163835	0.00693281	153.13243000	0.00000000
ϕ_1^*	-0.00772613	0.00322549	-2.39534000	0.01660499
ϕ_2^*	0.00749174	0.00322533	2.32279000	0.02019070
a_{11}^*	-0.12378718	0.02411502	-5.13320000	0.00000028
a_{21}^*	0.12880290	0.02645591	4.86859000	0.00000112
		GARCH(1.1)		
$lpha_{arepsilon}$	0.00000051	0.00000009	5.37402000	0.00000008
$oldsymbol{eta}_{1,arepsilon}$	0.05982310	0.01050350	5.69554000	0.0000001
$oldsymbol{eta}_{2,arepsilon}$	0.90478290	0.01472628	61.44000000	0.00000000

b)The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	3.2829	0.511639
LB(8)	12.4467	0.132359
LB(12)	18.6877	0.096348
LB(16)	20.2947	0.207241
LB(20)	21.9339	0.344106
LB(24)	23.1336	0.511921

	Test Statistic	Significance Level
ARCH(4)	0.82933	0.50634
ARCH(8)	1.57515	0.12705
ARCH(12)	1.57844	0.09090
ARCH(16)	1.24878	0.22234
ARCH(20)	1.07612	0.36796
ARCH(24)	0.95080	0.53074

Table I.6: Estimation of the conditional mean return and conditional volatility

Germany

a)Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
α_i^*	-0.00510716	0.00223699	-2.28305000	0.02242723
$oldsymbol{eta}_i^*$	1.05635938	0.01098466	96.16674000	0.00000000
ϕ_1^*	-0.00220653	0.00139170	-1.58549000	0.11285399
ϕ_2^*	0.00338574	0.00166885	2.02878000	0.04248052
a_{11}^*	-0.05396259	0.02687695	-2.00776000	0.04466827
a_{21}^*	0.05453255	0.03055587	1.78468000	0.07431274
		GARCH(1.1)		
$lpha_arepsilon$	0.00000086	0.00000029	2.98633000	0.00282344
$oldsymbol{eta}_{1,arepsilon}$	0.00000127	0.00000020	6.40171000	0.00000000
$oldsymbol{eta}_{2,arepsilon}$	0.14814245	0.01850584	8.00517000	0.00000000

b) The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	0.99390	0.91072
LB(8)	4.70020	0.78909
LB(12)	7.33020	0.83504
LB(16)	15.42120	0.49403
LB(20)	20.22740	0.44379
LB(24)	21.63970	0.60076

	Test Statistic	Significance Level
ARCH(4)	0.24290	0.91403
ARCH(8)	0.56294	0.80892
ARCH(12)	0.59432	0.84845
ARCH(16)	0.94838	0.51222
ARCH(20)	0.92997	0.54814
ARCH(24)	0.81157	0.72543

Table I.7: Estimation of the conditional mean return and conditional volatility

Greece

a)Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
$lpha_i^*$	-0.00149160	0.00606720	-0.24585000	0.80579530
$oldsymbol{eta}_i^*$	0.58040000	0.02000000	29.00457000	0.00000000
ϕ_1^*	-0.00168510	0.00169310	-0.99529000	0.31959535
ϕ_2^*	0.00214340	0.00286500	0.74812000	0.45438976
a_{11}^*	0.03890000	0.02140000	1.81741000	0.06915471
a_{21}^*	0.07920000	0.02320000	3.41642000	0.00063449
		GARCH(1.1)		
$lpha_arepsilon$	0.00000213	0.00000074	2.86752000	0.00413709
$oldsymbol{eta}_{ ext{l},arepsilon}$	0.09100000	0.01640000	5.55704000	0.00000003
$oldsymbol{eta}_{2,arepsilon}$	0.89810000	0.01780000	50.43491000	0.00000000

b)The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	1.14560	0.886972
LB(8)	1.94670	0.982603
LB(12)	2.43920	0.998374
LB(16)	11.99000	0.744671
LB(20)	15.54660	0.744323
LB(24)	16.04580	0.886416

	Test Statistic	Significance Level
ARCH(4)	0.27661	0.89321
ARCH(8)	0.22604	0.98629
ARCH(12)	0.18073	0.99909
ARCH(16)	0.73093	0.76428
ARCH(20)	0.70246	0.82733
ARCH(24)	0.59406	0.93999

Table I.8: Estimation of the conditional mean return and conditional volatility

Holland

a)Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
$lpha_i^*$	-0.00326882	0.00223329	-1.46368000	0.14328207
$\boldsymbol{\beta}_i^*$	1.01445467	0.01046889	96.90186000	0.00000000
ϕ_1^*	-0.00229601	0.00156692	-1.46530000	0.14283918
ϕ_2^*	0.00295345	0.00167595	1.76225000	0.07802640
a_{11}^*	-0.02834799	0.02275514	-1.24578000	0.21284381
a_{21}^*	0.04265660	0.02528771	1.68685000	0.09163199
		GARCH(1.1)		
$lpha_arepsilon$	0.00000074	0.00000014	5.12097000	0.00000030
$oldsymbol{eta}_{1,arepsilon}$	0.06577327	0.01007172	6.53049000	0.00000000
$oldsymbol{eta}_{2,arepsilon}$	0.90972542	0.01304535	69.73559000	0.00000000

b) The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	2.5942	0.62786
LB(8)	5.4617	0.70728
LB(12)	7.5211	0.82134
LB(16)	8.3952	0.93622
LB(20)	9.9656	0.96879
LB(24)	11.8580	0.98146

	Test Statistic	Significance Level
ARCH(4)	0.65203	0.62545
ARCH(8)	0.69795	0.69370
ARCH(12)	0.61800	0.82862
ARCH(16)	0.51451	0.94129
ARCH(20)	0.48298	0.97363
ARCH(24)	0.47411	0.98580

Table I.9: Estimation of the conditional mean return and conditional volatility

Ireland

a)Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
α_i^*	0,00425530	0,00428410	0,99329000	0,32056772
$oldsymbol{eta}_i^*$	0,72050000	0,02000000	36,08621000	0,00000000
ϕ_1^*	-0,00005093	0,00197820	-0,02575000	0,97946004
ϕ_2^*	-0,00089682	0,00244610	-0,36663000	0,71389358
a_{11}^*	0,02300000	0,02340000	0,98343000	0,32539356
a_{21}^*	0,08790000	0,02500000	3,51527000	0,00043931
		GARCH(1.1)		
$lpha_arepsilon$	0,00000286	0,00000073	3,91252000	0,00009134
$oldsymbol{eta}_{1,arepsilon}$	0,10160000	0,01680000	6,03300000	0,00000000
$oldsymbol{eta}_{2,arepsilon}$	0,88280000	0,01890000	46,79823000	0,00000000

b)The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	2,05830	0,72504
LB(8)	5,63810	0,68769
LB(12)	11,02800	0,52652
LB(16)	11,71890	0,76309
LB(20)	19,33420	0,50021
LB(24)	20,78360	0,65146

	Test Statistic	Significance Level
ARCH(4)	0.51316	0.72608
ARCH(8)	0.69798	0.69367
ARCH(12)	0.90410	0.54206
ARCH(16)	0.73634	0.75847
ARCH(20)	0.96306	0.50526
ARCH(24)	0.85399	0.66785

Table I.10: Estimation of the conditional mean return and conditional volatility

Italy

a) Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
α_i^*	0.00431680	0.00217330	1.98627000	0.04700354
$oldsymbol{eta}_i^*$	0.85660000	0.00944520	90.69043000	0.00000000
ϕ_1^*	-0.00164880	0.00159250	-1.03540000	0.30048237
ϕ_2^*	0.00064822	0.00150790	0.42987000	0.66728930
a_{11}^*	-0.00820200	0.02460000	-0.33318000	0.73899502
a_{21}^*	0.00851660	0.02330000	0.36601000	0.71435517
		GARCH(1.1)		
$lpha_arepsilon$	0.00000051	0.0000014	3.55426000	0.00037904
$oldsymbol{eta}_{1,arepsilon}$	0.06730000	0.01230000	5.48509000	0.00000004
$oldsymbol{eta}_{2,arepsilon}$	0.91680000	0.01510000	60.59405000	0.00000000

b) The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	2.05600	0.72547
LB(8)	3.78790	0.87573
LB(12)	11.84640	0.45809
LB(16)	13.14740	0.66195
LB(20)	17.31810	0.63223
LB(24)	22.62880	0.54179

	Test Statistic	Significance Level
ARCH(4)	0.51344	0.72587
ARCH(8)	0.46606	0.88059
ARCH(12)	0.96535	0.47997
ARCH(16)	0.78063	0.70930
ARCH(20)	0.84209	0.66293
ARCH(24)	0.89974	0.60338

Table I.11: Estimation of the conditional mean return and conditional volatility

Norway

a) Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
$lpha_i^*$	-0,00158160	0,00489140	-0,32335000	0,74643383
$oldsymbol{eta}_i^*$	0,72240000	0,02260000	31,95025000	0,00000000
ϕ_1^*	-0,00054342	0,00085510	-0,63550000	0,52509976
ϕ_2^*	0,00112790	0,00166580	0,67708000	0,49835206
a_{11}^*	-0,04000000	0,02240000	-1,78293000	0,07459801
a_{21}^*	0,16130000	0,02590000	6,23939000	0,00000000
		GARCH(1.1)		
$lpha_arepsilon$	0,00000330	0,00000092	3,57426000	0,00035122
$oldsymbol{eta}_{1,arepsilon}$	0,08650000	0,01320000	6,54817000	0,00000000
$oldsymbol{eta}_{2,arepsilon}$	0,89480000	0,01650000	54,37460000	0,00000000

b)The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	1,17220	0,88265
LB(8)	4,73520	0,78547
LB(12)	5,47820	0,94007
LB(16)	8,11260	0,94544
LB(20)	13,76690	0,84211
LB(24)	16,40680	0,87284

	Test Statistic	Significance Level
ARCH(4)	0.29091	0.88401
ARCH(8)	0.60072	0.77798
ARCH(12)	0.47103	0.93232
ARCH(16)	0.53773	0.92843
ARCH(20)	0.70383	0.82593
ARCH(24)	0.70734	0.84883

Table I.12: Estimation of the conditional mean return and conditional volatility

Portugal

a) Coefficients of the conditional mean and conditional volatility

Q ₀ o F F	Tatimata	O+ d - 11	П Оьоь	۵ : ۵۰۰۰ : ۶
Coeff	Estimate	Std Error	T-Stat	Signif
$lpha_i^*$	-0.00326190	0.00290070	-1.12451000	0.26079871
$oldsymbol{eta}_i^*$	0.49900000	0.01320000	37.74321000	0.00000000
ϕ_1^*	-0.00760090	0.00241120	-3.15228000	0.00162002
ϕ_2^*	0.00848250	0.00258540	3.28088000	0.00103486
a_{11}^*	0.06280000	0.02160000	2.90420000	0.00368187
a_{21}^*	0.02220000	0.01680000	1.32216000	0.18611529
		GARCH(1.1)		
$\alpha_{arepsilon}$	0.00000051	0.00000022	2.32173000	0.02024730
$oldsymbol{eta}_{1,arepsilon}$	0.05880000	0.01060000	5.54828000	0.00000003
$oldsymbol{eta}_{2,arepsilon}$	0.93520000	0.01210000	76.99398000	0.00000000

b) The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	3.86790	0.42419
LB(8)	10.17830	0.25274
LB(12)	13.16710	0.35701
LB(16)	24.28640	0.08343
LB(20)	32.60080	0.03730
LB(24)	34.70190	0.07295

	Test Statistic	Significance Level
ARCH(4)	0.94552	0.43652
ARCH(8)	1.33206	0.22262
ARCH(12)	1.12704	0.33284
ARCH(16)	1.45472	0.10782
ARCH(20)	1.52030	0.06484
ARCH(24)	1.36660	0.11019

Table I.13: Estimation of the conditional mean return and conditional volatility

Spain

a) Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
$lpha_i^*$	0.00116230	0.00239600	0.48509000	0.62761464
$\boldsymbol{\beta}_i^*$	0.95980000	0.01100000	86.94178000	0.00000000
ϕ_1^*	-0.00111340	0.00102760	-1.08347000	0.27860133
ϕ_2^*	0.00096242	0.00127760	0.75330000	0.45127104
a_{11}^*	0.04890000	0.02590000	1.88592000	0.05930604
a_{21}^*	-0.05720000	0.02650000	-2.15454000	0.03119793
		GARCH(1.1)		
$lpha_{arepsilon}$	0.00000045	0.00000012	3.65950000	0.00025270
$oldsymbol{eta}_{ ext{l},arepsilon}$	0.05730000	0.01050000	5.47143000	0.00000004
$oldsymbol{eta}_{2,arepsilon}$	0.93200000	0.01190000	78.52230000	0.00000000

b) The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	3.90480	0.41904
LB(8)	10.08120	0.25937
LB(12)	15.24790	0.22816
LB(16)	19.01110	0.26809
LB(20)	26.76900	0.14189
LB(24)	28.77460	0.22873

	Test Statistic	Significance Level
ARCH(4)	0.96857	0.42351
ARCH(8)	1.26886	0.25525
ARCH(12)	1.27292	0.22770
ARCH(16)	1.15163	0.30058
ARCH(20)	1.28409	0.17815
ARCH(24)	1.14269	0.28660

Table I.14: Estimation of the conditional mean return and conditional volatility

Sweden

a) Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
*	Escillace	DCG BITOI	1 Deac	Bigiii
α_i^*	-0.00121538	0.00493473	-0.24629000	0.80545686
$oldsymbol{eta}_i^*$	1.11693257	0.01782879	62.64770000	0.00000000
ϕ_1^*	-0.00507551	0.00223429	-2.27165000	0.02310792
ϕ_2^*	0.00544122	0.00311505	1.74675000	0.08067997
$a_{11,1}^{*}$	-0.03131529	0.02209709	-1.41717000	0.15643382
$a_{11,2}^{*}$	-0.02107185	0.01273871	-1.65416000	0.09809527
a_{21}^*	0.10050070	0.02922101	3.43933000	0.00058316
		GARCH(1.1)		
$lpha_arepsilon$	0.0000065	0.00000027	2.35483000	0.01853125
$oldsymbol{eta}_{1,arepsilon}$	0.03230886	0.00850062	3.80077000	0.00014425
$eta_{2,arepsilon}$	0.96046012	0.01066594	90.04931000	0.00000000

b) The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	6.6014	0.15851
LB(8)	13.8824	0.08488
LB(12)	17.8977	0.11883
LB(16)	21.7491	0.15148
LB(20)	27.2499	0.12838
LB(24)	29.0949	0.21658

	Test Statistic	Significance Level
ARCH(4)	1.65831	0.15700
ARCH(8)	1.90722	0.05490
ARCH(12)	1.65973	0.06955
ARCH(16)	1.43682	0.11530
ARCH(20)	1.43213	0.09652
ARCH(24)	1.22683	0.20568

Table I.15: Estimation of the conditional mean return and conditional volatility

Switzerland

a) Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif
α_i^*	0,00576680	0,00253750	2,27267000	0,02304619
$oldsymbol{eta}_i^*$	0,75950000	0,01090000	69,69342000	0,00000000
ϕ_1^*	-0,00684150	0,00217280	-3,14867000	0,00164013
ϕ_2^*	0,00564370	0,00202370	2,78884000	0,00528964
a_{11}^*	-0,07540000	0,02300000	-3,28208000	0,00103043
a_{21}^*	0,09690000	0,02100000	4,62025000	0,00000383
		GARCH(1.1)		
$lpha_arepsilon$	0,00000086	0,00000029	2,98633000	0,00282344
$oldsymbol{eta}_{1,arepsilon}$	0,08190000	0,01620000	5,05588000	0,00000043
$oldsymbol{eta}_{2,arepsilon}$	0,89610000	0,02140000	41,96684000	0,00000000

b) The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	7,0595	0,13278
LB(8)	12,9620	0,11317
LB(12)	16,2472	0,18017
LB(16)	20,9701	0,17966
LB(20)	23,3852	0,27028
LB(24)	26,0011	0,35311

	Test Statistic	Significance Level
ARCH(4)	1.73591	0.13931
ARCH(8)	1.58688	0.12351
ARCH(12)	1.32169	0.19872
ARCH(16)	1.22135	0.24278
ARCH(20)	1.04627	0.40237
ARCH(24)	0.97108	0.50221

Table I.16: Estimation of the conditional mean return and conditional volatility

United Kingdom

a) Coefficients of the conditional mean and conditional volatility

Coeff	Estimate	Std Error	T-Stat	Signif	
$lpha_i^*$	0.00288720	0.00231280	1.24837000	0.21189393	
$oldsymbol{eta}_i^*$	0.98020000	0.00841430	116.49217000	0.00000000	
ϕ_1^*	-0.00230410	0.00192380	-1.19768000	0.23104055	
ϕ_2^*	0.00162940	0.00160230	1.01692000	0.30919275	
a_{11}^*	-0.11170000	0.02560000	-4.36337000	0.00001281	
a_{21}^*	0.08310000	0.02650000	3.14186000	0.00167879	
	GARCH(1.1)				
$lpha_arepsilon$	0.00000026	0.00000009	2.78198000	0.00540280	
$oldsymbol{eta}_{ ext{l},arepsilon}$	0.07090000	0.01590000	4.47269000	0.00000772	
$oldsymbol{eta}_{2,arepsilon}$	0.91980000	0.01720000	53.52706000	0.00000000	

b) The Ljung-Box Qui- Squared Test for Serial Correlation in Normalized Residuals (number of lags within parenthesis)

	Test Statistic	Significance Level
LB(4)	6.19150	0.18530
LB(8)	9.40680	0.30915
LB(12)	13.83480	0.31138
LB(16)	22.18810	0.13722
LB(20)	24.63580	0.21572
LB(24)	29.06040	0.21787

	Test Statistic	Significance Level
ARCH(4)	1.50043	0.19942
ARCH(8)	1.08780	0.36849
ARCH(12)	1.02096	0.42619
ARCH(16)	1.24674	0.22382
ARCH(20)	1.13522	0.30478
ARCH(24)	1.12395	0.30708

Table II Statistics on the Sharpe ratios

Statistics on	the Condit	ional Sharp	e Ratio	Ex Post Sharpe Ratio
		AUS'	TRIA	
SUB-SAMPLE		STD ERROR	SIG LEVEL	
2001-2009		0.65741	0.00001	-0.00057
2001-2002	0.01826	0.82904	0.61642	-0.01780
2003-2004	0.13459	0.65486	0.00000	0.15236
2005-2006	0.11664	0.49503	0.00000	0.09525
2007-2009	-0.00830	0.60436	0.73086	-0.06620
Test for equality as	cross the sub	samples:		
Chi-Squared(4)=	13.675877 wi	th Significano	ce Level 0.00840471	
			GIUM	
SUB-SAMPLE	MEAN	STD ERROR	SIG LEVEL	
2001-2009	0.04161	1.37614	0.15691	-0.02934
2001-2002	-0.08182	1.45462	0.20106	-0.05760
2003-2004	0.16143	1.30841	0.00513	0.05153
2005-2006	0.14597	1.32747	0.01246	0.08493
2007-2009	-0.04782	1.39329	0.39008	-0.08765
Test for equality a				
Chi-Squared(4)=			ce Level 0.00000008	
			MARK	1
SUB-SAMPLE	MEAN	STD ERROR	SIG LEVEL	
2001-2009		0.93667	0.23355	-0.00252
	-0.04979	1.00341	0.25923	-0.00232
2001-2002	0.07672	0.90266	0.25923	0.06995
2005-2004	0.12511	0.69497	0.00005	0.00393
2007-2009	-0.04758	1.06525	0.26346	-0.03351
			0.20340	-0.03351
Test for equality as				
Chi-Squared(4)=	18.384682 W1			
			LAND	
SUB-SAMPLE		STD ERROR	SIG LEVEL	
	-0.03852	1.19238	0.13049	-0.02519
2001-2002	-0.11278	0.85449	0.00279	-0.05277
2003-2004	-0.01208	0.77649	0.72316	-0.01385
2005-2006	0.13342	0.97435	0.00189	0.05819
2007-2009	-0.14601	1.74170	0.03605	-0.03560
Test for equality a				
Chi-Squared(4)=	31.957941 wi	th Significano	ce Level 0.00000195	
		FRA	NCE	
SUB-SAMPLE		STD ERROR	SIG LEVEL	
2001-2009	-0.05269	3.62067	0.49560	-0.02079
2001-2002	-0.24666	3.63305	0.12290	-0.06591
2003-2004	0.15747	2.83163	0.20577	0.02334
2005-2006	0.20669	2.77706	0.09026	0.07833
2007-2009	-0.29340	4.65362	0.11461	-0.04736
Test for equality ac	cross the sub	samples:		
Chi-Squared(4)=	177.228993 wi	th Significano	e Level 0.00000000	
 		GER	MANY	•
SUB-SAMPLE	MEAN	STD ERROR	SIG LEVEL	
2001-2009	-0.02929	2.29488	0.55013	-0.01750
2001-2002	-0.21987	2.02937	0.01399	-0.07744
2001-2002	0.06201	1.63600	0.38823	0.03197
2005-2004	0.06201	2.09466	0.00542	0.03197
2007-2009	-0.18914	3.00314	0.11500	-0.04063
Test for equality a			0.11300	-0.04063
Chi-Squared(4)=			ce Level 0.00000000	
CIII-Dduaren(4)=	70.740300 WI		ECE	1
		GRE	11-C11-	1
OIID CAMPIT	1.4T3 73 3 T	CMD HDDOD	CIC I EXTER	
SUB-SAMPLE		STD ERROR	SIG LEVEL	0.0000
2001-2009	0.02100	0.63696	0.12276	-0.02004
2001-2002	-0.01445	0.76038	0.66556	-0.10970
2003-2004	0.06599	0.56957	0.00855	0.09052
	0 07122	0.46660	0.00054	0.07690
2005-2006	0.07123	0.10000		
2005-2006 2007-2009	-0.03276	0.69210	0.23602	-0.05761
	-0.03276	0.69210		

Table II (Cont.)

the Condit	ional Sharp	e Ratio	Ex Post Sharpe
			Ratio
	HOL	LAND	
MEAN	STD ERROR	SIG LEVEL	-0.02696
			-0.06610
			-0.00258
	1.99905		0.08137
0.18879	1.89466	0.02348	-0.05270
-0.13985	2.65073	0.18660	
ross the sub	samples:		
03.285507 wi	th Significand	ce Level 0.00000000	
	IRE	LAND	
MEAN	STD ERROR	SIG LEVEL	
-0.00260	0.80509	0.87998	-0.03749
-0.09786	1.00975	0.02784	-0.06677
0.05263	0.78326	0.12645	0.06489
0.08096	0.70307	0.00890	0.05615
			-0.08985
		ce Level 0.00178993	
_,,,,,,,, WI			1
MFAN			
			-0.02853
			-0.02853
			0.04635
			0.06010
		0.09383	-0.06485
ross the sub	samples:	- 1 0 0000000	
66.133349 W1			
			T.
			-0.00109
			-0.05524
0.08184	0.72127		0.07167
0.10316	0.45790	0.00000	0.06745
-0.00082	0.69154	0.97616	-0.03123
9.938038 wi			
	PORT	CUGAL	
MEAN	STD ERROR	SIG LEVEL	
0 00010	0.78909	0.58599	
0.00918		0.56599	-0.02596
-0.01085	0.87996	0.77913	-0.02596 -0.09535
-0.01085	0.87996	0.77913	-0.09535
-0.01085 0.01373	0.87996 0.65251	0.77913 0.63180	-0.09535 0.05197
-0.01085 0.01373 0.12719 -0.07961	0.87996 0.65251 0.68606 0.87795	0.77913 0.63180 0.00003	-0.09535 0.05197 0.09672
-0.01085 0.01373 0.12719 -0.07961 cross the sub	0.87996 0.65251 0.68606 0.87795 samples:	0.77913 0.63180 0.00003 0.02340	-0.09535 0.05197 0.09672
-0.01085 0.01373 0.12719 -0.07961 cross the sub	0.87996 0.65251 0.68606 0.87795 samples: th Significanc	0.77913 0.63180 0.00003	-0.09535 0.05197 0.09672
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi	0.87996 0.65251 0.68606 0.87795 samples: th Significand	0.77913 0.63180 0.00003 0.02340 ce Level 0.00250519 AIN	-0.09535 0.05197 0.09672
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi	0.87996 0.65251 0.68606 0.87795 samples: th Significand SPA	0.77913 0.63180 0.00003 0.02340 Ce Level 0.00250519 AIN	-0.09535 0.05197 0.09672 -0.06269
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi	0.87996 0.65251 0.68606 0.87795 samples: th Significance SPI STD ERROR 1.93966	0.77913 0.63180 0.00003 0.02340 Ce Level 0.00250519 AIN SIG LEVEL 0.81096	-0.09535 0.05197 0.09672 -0.06269
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi MEAN 0.00991 -0.10571	0.87996 0.65251 0.68606 0.87795 samples: th Significance SPI STD ERROR 1.93966 1.75071	0.77913 0.63180 0.00003 0.02340 Ce Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997	-0.09535 0.05197 0.09672 -0.06269 -0.00566 -0.05210
-0.01085 0.01373 0.12719 -0.07961 ross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323	0.87996 0.65251 0.68606 0.87795 samples: th Significance SPI STD ERROR 1.93966 1.75071 1.77975	0.77913 0.63180 0.00003 0.02340 Ce Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783	-0.09535 0.05197 0.09672 -0.06269 -0.00566 -0.05210 0.05396
-0.01085 0.01373 0.12719 -0.07961 ross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583	0.87996 0.65251 0.68606 0.87795 samples: th Significand SPI STD ERROR 1.93966 1.75071 1.77975 1.84095	0.77913 0.63180 0.00003 0.02340 SEE Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783 0.02174	-0.09535 0.05197 0.09672 -0.06269 -0.00566 -0.05210 0.05396 0.10293
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583 -0.13306	0.87996 0.65251 0.68606 0.87795 samples: th Significand SPI STD ERROR 1.93966 1.75071 1.77975 1.84095 2.25869	0.77913 0.63180 0.00003 0.02340 Ce Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783	-0.09535 0.05197 0.09672 -0.06269 -0.00566 -0.05210 0.05396
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583 -0.13306 cross the sub	0.87996 0.65251 0.68606 0.87795 samples: th Significand SPI STD ERROR 1.93966 1.75071 1.77975 1.84095 2.25869 samples:	0.77913 0.63180 0.00003 0.02340 SEE Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783 0.02174 0.14037	-0.09535 0.05197 0.09672 -0.06269 -0.00566 -0.05210 0.05396 0.10293
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583 -0.13306 cross the sub	0.87996 0.65251 0.68606 0.87795 samples: th Significand SPI STD ERROR 1.93966 1.75071 1.77975 1.84095 2.25869 samples: th Significand	0.77913 0.63180 0.00003 0.02340 SEE Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783 0.02174 0.14037 SEE Level 0.00000000	-0.09535 0.05197 0.09672 -0.06269 -0.00566 -0.05210 0.05396 0.10293
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583 -0.13306 cross the sub 63.853824 wi	0.87996 0.65251 0.68606 0.87795 samples: th Significand SPI STD ERROR 1.93966 1.75071 1.77975 1.84095 2.25869 samples: th Significand SWE	0.77913 0.63180 0.00003 0.02340 SEE Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783 0.02174 0.14037 SEE Level 0.00000000	-0.09535 0.05197 0.09672 -0.06269 -0.0566 -0.05210 0.05396 0.10293
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583 -0.13306 cross the sub 63.853824 wi	0.87996 0.65251 0.68606 0.87795 samples: th Significance SPI STD ERROR 1.93966 1.75071 1.77975 1.84095 2.25869 samples: th Significance SWE STD ERROR	0.77913 0.63180 0.00003 0.02340 CE Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783 0.02174 0.14037 CE Level 0.00000000 CDEN SIG LEVEL	-0.09535 0.05197 0.09672 -0.06269 -0.05266 -0.05210 0.05396 0.10293 -0.03711
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583 -0.13306 cross the sub 63.853824 wi	0.87996 0.65251 0.68606 0.87795 samples: th Significance SPE STD ERROR 1.93966 1.75071 1.77975 1.84095 2.25869 samples: th Significance SWE STD ERROR	0.77913 0.63180 0.00003 0.02340 Ce Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783 0.02174 0.14037 Ce Level 0.000000000 CDEN SIG LEVEL 0.97399	-0.09535 0.05197 0.09672 -0.06269 -0.05266 -0.05210 0.05396 0.10293 -0.03711
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583 -0.13306 cross the sub 63.853824 wi MEAN 0.00104 -0.02120	0.87996 0.65251 0.68606 0.87795 samples: th Significanc SPE STD ERROR 1.93966 1.75071 1.77975 1.84095 2.25869 samples: th Significanc SWE STD ERROR 1.49984 1.35527	0.77913 0.63180 0.00003 0.02340 Ce Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783 0.02174 0.14037 Ce Level 0.00000000 CDEN SIG LEVEL 0.97399 0.72197	-0.09535 0.05197 0.09672 -0.06269 -0.00566 -0.05210 0.05396 0.10293 -0.03711
-0.01085 0.01373 0.12719 -0.07961 ross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583 -0.13306 ross the sub 63.853824 wi MEAN 0.00104 -0.02120 0.09744	0.87996 0.65251 0.68606 0.87795 samples: th Significance STD ERROR 1.93966 1.75071 1.77975 1.84095 2.25869 samples: th Significance SWE STD ERROR 1.49984 1.35527 1.30724	0.77913 0.63180 0.00003 0.02340 Ce Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783 0.02174 0.14037 Ce Level 0.000000000 CDEN SIG LEVEL 0.97399	-0.09535 0.05197 0.09672 -0.06269 -0.05266 -0.05210 0.05396 0.10293 -0.03711
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583 -0.13306 cross the sub 63.853824 wi MEAN 0.00104 -0.02120 0.09744 0.09306	0.87996 0.65251 0.68606 0.87795 samples: th Significance STD ERROR 1.93966 1.75071 1.77975 1.84095 2.25869 samples: th Significance SWE STD ERROR 1.49984 1.35527 1.30724 1.17084	0.77913 0.63180 0.00003 0.02340 Ce Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783 0.02174 0.14037 Ce Level 0.00000000 CDEN SIG LEVEL 0.97399 0.72197	-0.09535 0.05197 0.09672 -0.06269 -0.00566 -0.05210 0.05396 0.10293 -0.03711
-0.01085 0.01373 0.12719 -0.07961 ross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583 -0.13306 ross the sub 63.853824 wi MEAN 0.00104 -0.02120 0.09744	0.87996 0.65251 0.68606 0.87795 samples: th Significance STD ERROR 1.93966 1.75071 1.77975 1.84095 2.25869 samples: th Significance SWE STD ERROR 1.49984 1.35527 1.30724	0.77913 0.63180 0.00003 0.02340 Ce Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783 0.02174 0.14037 Ce Level 0.00000000 CDEN SIG LEVEL 0.97399 0.72197 0.09008	-0.09535 0.05197 0.09672 -0.06269 -0.00566 -0.05210 0.05396 0.10293 -0.03711
-0.01085 0.01373 0.12719 -0.07961 cross the sub 16.419283 wi MEAN 0.00991 -0.10571 0.11323 0.18583 -0.13306 cross the sub 63.853824 wi MEAN 0.00104 -0.02120 0.09744 0.09306	0.87996 0.65251 0.68606 0.87795 samples: th Significance STD ERROR 1.93966 1.75071 1.77975 1.84095 2.25869 samples: th Significance SWE STD ERROR 1.49984 1.35527 1.30724 1.17084 1.93519	0.77913 0.63180 0.00003 0.02340 SEE Level 0.00250519 AIN SIG LEVEL 0.81096 0.16997 0.14783 0.02174 0.14037 SEE Level 0.00000000 SDEN SIG LEVEL 0.97399 0.72197 0.09008 0.07048	-0.09535 0.05197 0.09672 -0.06269 -0.00566 -0.05210 0.05396 0.10293 -0.03711 -0.01487 -0.06319 0.06066 0.07017
	-0.05062 -0.26361 0.01752 0.18879 -0.13985 ross the sub 03.285507 wi MEAN -0.00260 -0.09786 0.05263 0.08096 -0.04395 ross the sub 17.171741 wi MEAN -0.04409 -0.17488 0.06417 0.10045 -0.15432 ross the sub 66.133349 wi MEAN 0.04061 -0.01872 0.08184 0.10316 -0.00082 ross the sub 9.938038 wi	MEAN STD ERROR -0.05062 2.27754 -0.26361 2.37994 0.01752 1.99905 0.18879 1.89466 -0.13985 2.65073 ross the subsamples: 03.285507 with Significanc IRE MEAN STD ERROR -0.00260 0.80509 -0.09786 1.00975 0.05263 0.78326 0.08096 0.70307 -0.04395 0.69418 ross the subsamples: 17.171741 with Significanc IT MEAN STD ERROR -0.04409 1.97730 -0.17488 2.00531 0.06417 1.80998 0.10045 1.64125 -0.15432 2.30454 ross the subsamples: 66.133349 with Significanc NOF MEAN STD ERROR 0.04061 0.75575 -0.01872 1.0	-0.05062 2.27754 0.29804 -0.26361 2.37994 0.01200 0.01752 1.99905 0.84186 0.18879 1.89466 0.02348 -0.13985 2.65073 0.18660 ross the subsamples: 03.285507 with Significance Level 0.00000000

Table II (Cont.)

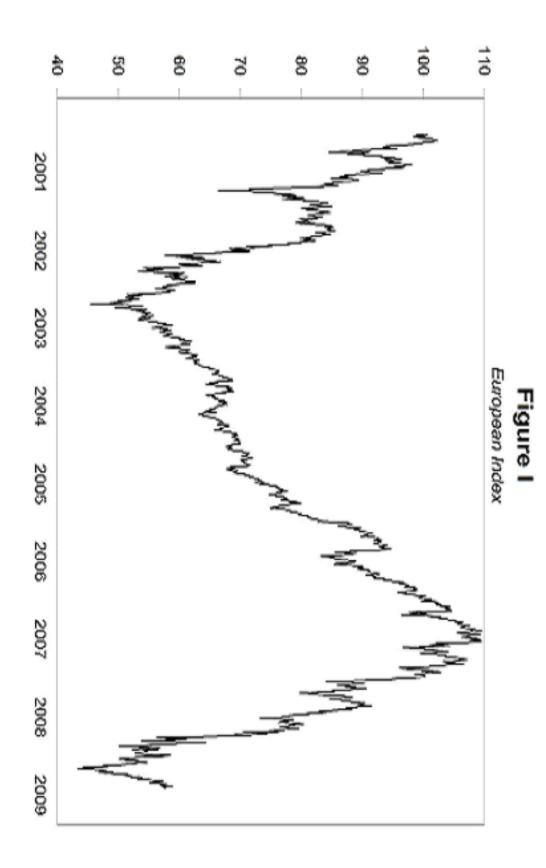
Statistics on	the Condit	ional Sharp	e Ratio	Ex Post Sharpe
				Ratio
		SWITZ	ERLAND	
SUB-SAMPLE MEAN	STD ERR	OR SIG LE	VEL	
2001-2009	-0.02269	1.57521	0.50012	-0.01979
2001-2002	-0.12652	1.64614	0.08084	-0.06672
2003-2004	0.10948	1.32491	0.06034	0.01319
2005-2006	0.07163	1.23927	0.18807	0.10564
2007-2009	-0.13252	1.90896	0.08241	-0.05500
Test for equality ac	ross the sub	samples:		
Chi-Squared(4)=	46.794663 wi	th Significano	e Level 0.00000000	
		UNITED	KINGDOM	
SUB-SAMPLE	MEAN	STD ERROR	SIG LEVEL	
2001-2009	-0.06877	2.77699	0.24627	-0.02548
2001-2002	-0.26357	2.93487	0.04146	-0.06066
2003-2004	0.08375	2.48447	0.44288	0.01602
2005-2006	0.11496	2.12876	0.21872	0.06513
2007-2009	-0.19824	3.28943	0.13149	-0.05242
Test for equality ac	ross the sub	samples:		
Chi-Squared(4)= 1	17.202131 wi	th Significano	e Level 0.00000000	

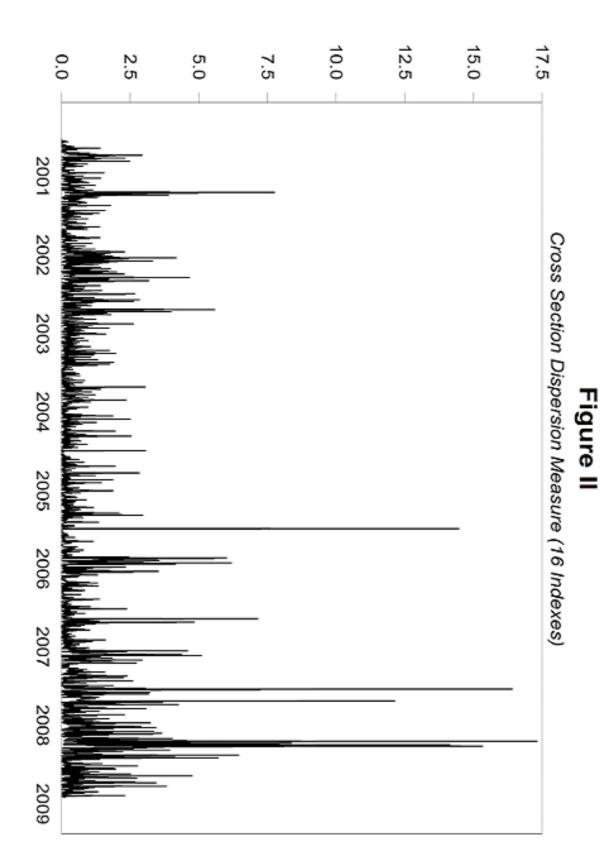
Table III – Statistics on the Cross Section Dispersion Measure between the Conditional Sharpe Ratios of the 16 stock indexes

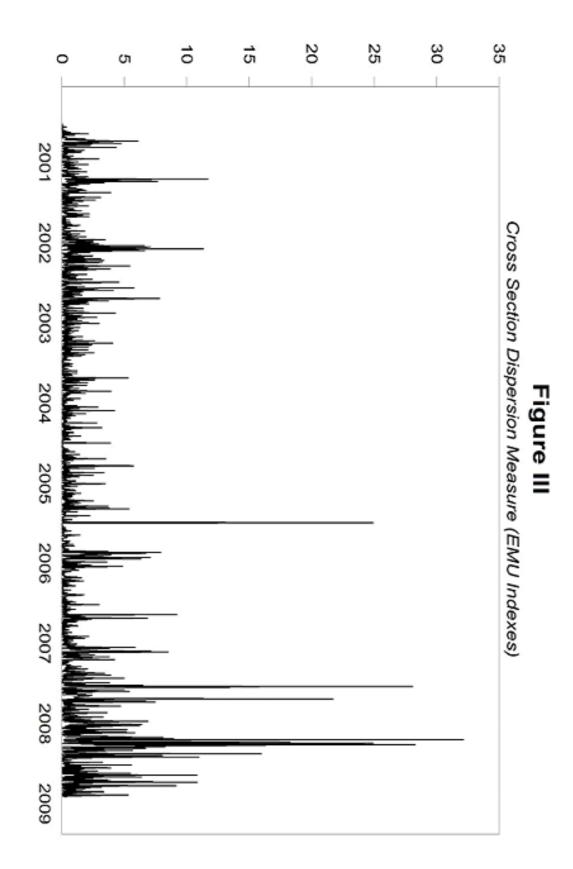
SUB-SAMPLE	MEAN	STD ERROR	SIG LEVEL	
2001-2009	0.80911	1.87144	0.0000	
2001-2002	0.74513	1.20997	0.0000	
2003-2004	0.50410	0.81653	0.0000	
2005-2006	0.51517	1.33881	0.0000	
2007-2009	1.36118	2.92020	0.0000	
Test for equality acr	coss the sub	samples:		
Chi-Squared(4)= 29	5.674687 wi	th Significan	ce Level 0.0000000	

Table IV – Statistics on the Cross Section Dispersion Measure between the Conditional Sharpe Ratios of the 11 EMU members stock indexes

SUB-SAMPLE	MEAN	STD ERROR	SIG LEVEL	
2001-2009	0.90064	2.11351	0.0000	
2001-2002	0.81303	1.35733	0.0000	
2003-2004	0.54376	0.87134	0.0000	
2005-2006	0.55471	1.48043	0.0000	
2007-2009	1.55949	3.31590	0.0000	
Test for equality a	cross the sub	samples:		
Chi-Squared(4)=	414.279744 wi	th Significan	ce Level 0.0000000	







ESTUDOS DO G.E.M.F.

(Available on-line at http://gemf.fe.uc.pt)

- 2009-16 The performance of the European Stock Markets: a time-varying Sharpe ratio approach José A. Soares da Fonseca
- 2009-15 Exchange Rate Mean Reversion within a Target Zone: Evidence from a Country on the Periphery of the ERM
 - António Portugal Duarte, João Sousa Andrade & Adelaide Duarte
- 2009-14 The Extent of Collective Bargaining and Workplace Representation: Transitions between States and their Determinants. A Comparative Analysis of Germany and Great Britain John T. Addison, Alex Bryson, Paulino Teixeira, André Pahnke & Lutz Bellmann
- 2009-13 How well the balance-of- payments constraint approach explains the Portuguese growth performance. Empirical evidence for the 1965-2008 period
 Micaela Antunes & Elias Soukiazis
- 2009-12 Atypical Work: Who Gets It, and Where Does It Lead? Some U.S. Evidence Using the NLSY79
 - John T. Addison, Chad Cotti & Christopher J. Surfield
- 2009-11 The PIGS, does the Group Exist? An empirical macroeconomic analysis based on the Okun Law
 - João Sousa Andrade
- 2009-10 A Política Monetária do BCE. Uma estratégia original para a estabilidade nominal João Sousa Andrade
- 2009-09 Wage Dispersion in a Partially Unionized Labor Force
 - John T. Addison, Ralph W. Bailey & W. Stanley Siebert
- 2009-08 Employment and exchange rates: the role of openness and technology Fernando Alexandre, Pedro Bação, João Cerejeira & Miguel Portela
- 2009-07 Channels of transmission of inequality to growth: A survey of the theory and evidence from a Portuguese perspective
 - Adelaide Duarte & Marta Simões
- 2009-06 No Deep Pockets: Some stylized results on firms' financial constraints
 - Filipe Silva & Carlos Carreira
- $2009 \hbox{-} 05$ Aggregate and sector-specific exchange rate indexes for the Portuguese economy
 - Fernando Alexandre, Pedro Bação, João Cerejeira & Miguel Portela
- 2009-04 Rent Seeking at Plant Level: An Application of the Card-De La Rica Tenure Model to Workers in German Works Councils
 - John T. Addison, Paulino Teixeira & Thomas Zwick
- 2009-03 Unobserved Worker Ability, Firm Heterogeneity, and the Returns to Schooling and Training
 Ana Sofia Lopes & Paulino Teixeira
- 2009-02 Worker Directors: A German Product that Didn't Export?
 - John T. Addison & Claus Schnabel
- 2009-01 Fiscal and Monetary Policies in a Keynesian Stock-flow Consistent Model
 - Edwin Le Heron
- 2008-08 Uniform Price Market and Behaviour Pattern: What does the Iberian Electricity Market Point Out
 - Vítor Marques, Isabel Soares & Adelino Fortunato
- 2008-07 The partial adjustment factors of FTSE 100 stock index and stock index futures: The informational impact of electronic trading systems
 - Helder M. C. V. Sebastião
- 2008-06 Water Losses and Hydrographical Regions Influence on the Cost Structure of the Portuguese Water Industry
 - Rita Martins, Fernando Coelho& Adelino Fortunato

- 2008-05 The Shadow of Death: Analysing the Pre-Exit Productivity of Portuguese Manufacturing Firms
 - Carlos Carreira & Paulino Teixeira
- 2008-04 A Note on the Determinants and Consequences of Outsourcing Using German Data
 John T. Addison, Lutz Bellmann, André Pahnke & Paulino Teixeira
- 2008-03 Exchange Rate and Interest Rate Volatility in a Target Zone: The Portuguese Case
 - António Portugal Duarte, João Sousa Andrade & Adelaide Duarte
- 2008-02 Taylor-type rules versus optimal policy in a Markov-switching economy
 - Fernando Alexandre, Pedro Bação & Vasco Gabriel
- 2008-01 Entry and exit as a source of aggregate productivity growth in two alternative technological regimes
 - Carlos Carreira & Paulino Teixeira
- 2007-09 Optimal monetary policy with a regime-switching exchange rate in a forward-looking model
 - Fernando Alexandre, Pedro Bação & John Driffill
- 2007-08 Estrutura económica, intensidade energética e emissões de CO₂: Uma abordagem Input-Output
 - Luís Cruz & Eduardo Barata
- 2007-07 The Stability and Growth Pact, Fiscal Policy Institutions, and Stabilization in Europe
 - Carlos Fonseca Marinheiro
- 2007-06 The Consumption-Wealth Ratio Under Asymmetric Adjustment
 - Vasco J. Gabriel, Fernando Alexandre & Pedro Bação
- 2007-05 European Integration and External Sustainability of the European Union An application of the thesis of Feldstein-Horioka
 - João Sousa Andrade
- 2007-04 Uma Aplicação da Lei de Okun em Portugal
 - João Sousa Andrade
- 2007-03 Education and growth: an industry-level analysis of the Portuguese manufacturing sector
 - Marta Simões & Adelaide Duarte
- 2007-02 Levels of education, growth and policy complementarities
 - Marta Simões & Adelaide Duarte
- 2007-01 Internal and External Restructuring over the Cycle: A Firm-Based Analysis of Gross Flows and Productivity Growth in Portugal
 - Carlos Carreira & Paulino Teixeira
- 2006-09 Cost Structure of the Portuguese Water Industry: a Cubic Cost Function Application
 - Rita Martins, Adelino Fortunato & Fernando Coelho
- 2006-08 The Impact of Works Councils on Wages
 - John T. Addison, Paulino Teixeira & Thomas Zwick
- 2006-07 Ricardian Equivalence, Twin Deficits, and the Feldstein-Horioka puzzle in Egypt
 - Carlos Fonseca Marinheiro
- 2006-06 L'intégration des marchés financiers
 - José Soares da Fonseca
- 2006-05 The Integration of European Stock Markets and Market Timing
 - José Soares da Fonseca
- 2006-04 Mobilidade do Capital e Sustentabilidade Externa uma aplicação da tese de F-H a Portugal (1910-2004)
 - João Sousa Andrade
- 2006-03 Works Councils, Labor Productivity and Plant Heterogeneity: First Evidence from Quantile Regressions
 - Joachim Wagner, Thorsten Schank, Claus Schnabel & John T. Addison
- 2006-02 Does the Quality of Industrial Relations Matter for the Macroeconomy? A Cross-Country Analysis Using Strikes Data
 - John T. Addison & Paulino Teixeira

- 2006-01 Monte Carlo Estimation of Project Volatility for Real Options Analysis
 - Pedro Manuel Cortesão Godinho
- 2005-17 On the Stability of the Wealth Effect
 - Fernando Alexandre, Pedro Bação & Vasco J. Gabriel
- 2005-16 Building Blocks in the Economics of Mandates
 - John T. Addison, C. R. Barrett & W. S. Siebert
- 2005-15 Horizontal Differentiation and the survival of Train and Coach modes in medium range passenger transport, a welfare analysis comprising economies of scope and scale
 - Adelino Fortunato & Daniel Murta
- 2005-14 'Atypical Work' and Compensation
 - John T. Addison & Christopher J. Surfield
- 2005-13 The Demand for Labor: An Analysis Using Matched Employer-Employee Data from the German LIAB. Will the High Unskilled Worker Own-Wage Elasticity Please Stand Up?

 John T. Addison, Lutz Bellmann, Thorsten Schank & Paulino Teixeira
- 2005-12 Works Councils in the Production Process
 - John T. Addison, Thorsten Schank, Claus Schnabel & Joachim Wagnerd
- 2005-11 Second Order Filter Distribution Approximations for Financial Time Series with Extreme Outliers
 - J. Q. Smith & António A. F. Santos
- 2005-10 Firm Growth and Persistence of Chance: Evidence from Portuguese Microdata
 Blandina Oliveira & Adelino Fortunato
- 2005-09 Residential water demand under block rates a Portuguese case study
 Rita Martins & Adelino Fortunato
- 2005-08 Politico-Economic Causes of Labor Regulation in the United States: Alliances and Raising Rivals' Costs (and Sometimes Lowering One's Own)
 - John T. Addison
- 2005-07 Firm Growth and Liquidity Constraints: A Dynamic Analysis
 - Blandina Oliveira & Adelino Fortunato
- 2005-06 The Effect of Works Councils on Employment Change
 - John T. Addison & Paulino Teixeira
- 2005-05 Le Rôle de la Consommation Publique dans la Croissance: le cas de l'Union Européenne - João Sousa Andrade, Maria Adelaide Silva Duarte & Claude Berthomieu
- 2005-04 The Dynamics of the Growth of Firms: Evidence from the Services Sector
 - Blandina Oliveira & Adelino Fortunato
- 2005-03 The Determinants of Firm Performance: Unions, Works Councils, and Employee Involvement/High Performance Work Practices
 - John T. Addison
- 2005-02 Has the Stability and Growth Pact stabilised? Evidence from a panel of 12 European countries and some implications for the reform of the Pact
 - Carlos Fonseca Marinheiro
- 2005-01 Sustainability of Portuguese Fiscal Policy in Historical Perspective
 - Carlos Fonseca Marinheiro
- 2004-03 Human capital, mechanisms of technological diffusion and the role of technological shocks in the speed of diffusion. Evidence from a panel of Mediterranean countries
 - Maria Adelaide Duarte & Marta Simões

- 2004-02 What Have We Learned About The Employment Effects of Severance Pay? Further Iterations of Lazear et al.
 - John T. Addison & Paulino Teixeira
- 2004-01 How the Gold Standard Functioned in Portugal: an analysis of some macroeconomic aspects
 António Portugal Duarte & João Sousa Andrade
- 2003-07 Testing Gibrat's Law: Empirical Evidence from a Panel of Portuguese Manufacturing Firms
 Blandina Oliveira & Adelino Fortunato
- 2003-06 Régimes Monétaires et Théorie Quantitative du Produit Nominal au Portugal (1854-1998)
 João Sousa Andrade
- 2003-05 Causas do Atraso na Estabilização da Inflação: Abordagem Teórica e Empírica Vítor Castro
- 2003-04 The Effects of Households' and Firms' Borrowing Constraints on Economic Growth
 Maria da Conceição Costa Pereira
- 2003-03 Second Order Filter Distribution Approximations for Financial Time Series with Extreme Outliers
 - J. Q. Smith & António A. F. Santos
- 2003-02 Output Smoothing in EMU and OECD: Can We Forego Government Contribution? A risk sharing approach
 - Carlos Fonseca Marinheiro
- 2003-01 Um modelo VAR para uma Avaliação Macroeconómica de Efeitos da Integração Europeia da Economia Portuguesa
 - João Sousa Andrade
- 2002-08 Discrimination des facteurs potentiels de croissance et type de convergence de l'économie portugaise dans l'UE à travers la spécification de la fonction de production macro-économique. Une étude appliquée de données de panel et de séries temporelles Marta Simões & Maria Adelaide Duarte
- 2002-07 Privatisation in Portugal: employee owners or just happy employees?
 -Luís Moura Ramos & Rita Martins
- 2002-06 The Portuguese Money Market: An analysis of the daily session Fátima Teresa Sol Murta
- 2002-05 As teorias de ciclo políticos e o caso português
 - Rodrigo Martins
- 2002-04 Fundos de acções internacionais: uma avaliação de desempenho - Nuno M. Silva
- 2002-03 The consistency of optimal policy rules in stochastic rational expectations models
 David Backus & John Driffill
- 2002-02 The term structure of the spreads between Portuguese and German interest rates during stage II of EMU
 - José Soares da Fonseca
- 2002-01 O processo desinflacionista português: análise de alguns custos e benefícios António Portugal Duarte
- 2001-14 Equity prices and monetary policy: an overview with an exploratory model Fernando Alexandre & Pedro Bação

- 2001-13 A convergência das taxas de juro portuguesas para os níveis europeus durante a segunda metade da década de noventa
 - José Soares da Fonseca
- 2001-12 Le rôle de l'investissement dans l'éducation sur la croissance selon différentes spécifications du capital humain.
 - Adelaide Duarte & Marta Simões
- 2001-11 Ricardian Equivalence: An Empirical Application to the Portuguese Economy
 - Carlos Fonseca Marinheiro
- 2001-10 A Especificação da Função de Produção Macro-Económica em Estudos de Crescimento Económico.
 - Maria Adelaide Duarte e Marta Simões
- 2001-09 Eficácia da Análise Técnica no Mercado Accionista Português
 - Nuno Silva
- 2001-08 The Risk Premiums in the Portuguese Treasury Bills Interest Rates: Estimation by a cointegration method
 - José Soares da Fonseca
- 2001-07 Principais factores de crescimento da economia portuguesa no espaço europeu
 - Maria Adelaide Duarte e Marta Simões
- 2001-06 Inflation Targeting and Exchange Rate Co-ordination
 - Fernando Alexandre, John Driffill e Fabio Spagnolo
- 2001-05 Labour Market Transition in Portugal, Spain, and Poland: A Comparative Perspective
 - Paulino Teixeira
- 2001-04 Paridade do Poder de Compra e das Taxas de Juro: Um estudo aplicado a três países da UEM
 - António Portugal Duarte
- 2001-03 Technology, Employment and Wages
 - John T. Addison & Paulino Teixeira
- 2001-02 Human capital investment through education and economic growth. A panel data analysis based on a group of Latin American countries
 - Maria Adelaide Duarte & Marta Simões
- 2001-01 Risk Premiums in the Porutguese Treasury Bills Interest Rates from 1990 to 1998. An ARCH-M Approach
 - José Soares da Fonseca
- 2000-08 Identificação de Vectores de Cointegração: Análise de Alguns Exemplos
 - Pedro Miguel Avelino Bação
- 2000-07 Imunização e M-quadrado: Que relação?
 - Jorge Cunha
- 2000-06 Eficiência Informacional nos Futuros Lisbor 3M
 - Nuno M. Silva
- 2000-05 Estimation of Default Probabilities Using Incomplete Contracts Data
 - J. Santos Silva & J. Murteira
- 2000-04 Un Essaie d'Application de la Théorie Quantitative de la Monnaie à l'économie portugaise, 1854-1998
 - João Sousa Andrade

- 2000-03 Le Taux de Chômage Naturel comme un Indicateur de Politique Economique? Une application à l'économie portugaise
 - Adelaide Duarte & João Sousa Andrade
- 2000-02 La Convergence Réelle Selon la Théorie de la Croissance: Quelles Explications pour l'Union Européenne?
 - Marta Cristina Nunes Simões
- 2000-01 Política de Estabilização e Independência dos Bancos Centrais
 - João Sousa Andrade
- 1999-09 Nota sobre a Estimação de Vectores de Cointegração com os Programas CATS in RATS, PCFIML e EVIEWS
 - Pedro Miguel Avelino Bação
- 1999-08 A Abertura do Mercado de Telecomunicações Celulares ao Terceiro Operador: Uma Decisão Racional?
 - Carlos Carreira
- 1999-07 Is Portugal Really so Arteriosclerotic? Results from a Cross-Country Analysis of Labour Adjustment
 - John T. Addison & Paulino Teixeira
- 1999-06 The Effect of Dismissals Protection on Employment: More on a Vexed Theme
 John T. Addison, Paulino Teixeira e Jean-Luc Grosso
- 1999-05 A Cobertura Estática e Dinâmica através do Contrato de Futuros PSI-20. Estimação das Rácios e Eficácia Ex Post e Ex Ante
 - Helder Miguel C. V. Sebastião
- 1999-04 Mobilização de Poupança, Financiamento e Internacionalização de Carteiras
 - João Sousa Andrade
- 1999-03 Natural Resources and Environment
 - Adelaide Duarte
- 1999-02 L'Analyse Positive de la Politique Monétaire
 - Chistian Aubin
- 1999-01 Economias de Escala e de Gama nos Hospitais Públicos Portugueses: Uma Aplicação da Função de Custo Variável Translog
 - Carlos Carreira
- 1998-11 Equilíbrio Monetário no Longo e Curto Prazos Uma Aplicação à Economia Portuguesa João Sousa Andrade
- 1998-10 Algumas Observações Sobre o Método da Economia
 - João Sousa Andrade
- 1998-09 Mudança Tecnológica na Indústria Transformadora: Que Tipo de Viés Afinal?
 Paulino Teixeira
- 1998-08 Portfolio Insurance and Bond Management in a Vasicek's Term Structure of Interest Rates
 José Alberto Soares da Fonseca
- 1998-07 Financial Innovation and Money Demand in Portugal: A Preliminary Study
 - Pedro Miguel Avelino Bação
- 1998-06 The Stability Pact and Portuguese Fiscal Policy: the Application of a VAR Model
 - Carlos Fonseca Marinheiro
- 1998-05 A Moeda Única e o Processo de Difusão da Base Monetária
 - José Alberto Soares da Fonseca

1998-04	La Structure par Termes et la Volatilité des Taux d'intérêt LISBOR - José Alberto Soares da Fonseca
1998-03	Regras de Comportamento e Reformas Monetárias no Novo SMI - João Sousa Andrade
1998-02	Um Estudo da Flexibilidade dos Salários: o Caso Espanhol e Português - Adelaide Duarte e João Sousa Andrade
1998-01	Moeda Única e Internacionalização: Apresentação do Tema - João Sousa Andrade
1997-09	Inovação e Aplicações Financeiras em Portugal - Pedro Miguel Avelino Bação
1997-08	Estudo do Efeito Liquidez Aplicado à Economia Portuguesa - João Sousa Andrade
1997-07	An Introduction to Conditional Expectations and Stationarity - Rui Manuel de Almeida
1997-06	Definição de Moeda e Efeito Berlusconi - João Sousa Andrade
1997-05	A Estimação do Risco na Escolha dos Portafólios: Uma Visão Selectiva - António Alberto Ferreira dos Santos
1997-04	A Previsão Não Paramétrica de Taxas de Rentabilidade - Pedro Manuel Cortesão Godinho
1997-03	Propriedades Assimptóticas de Densidades - Rui Manuel de Almeida
1997-02	Co-Integration and VAR Analysis of the Term Structure of Interest Rates: an empirical study of the Portuguese money and bond markets -João Sousa Andrade & José Soares da Fonseca
1997-01	Repartição e Capitalização. Duas Modalidades Complementares de Financiamento das Reformas - Maria Clara Murteira
1996-08	A Crise e o Ressurgimento do Sistema Monetário Europeu - Luis Manuel de Aguiar Dias
1996-07	Housing Shortage and Housing Investment in Portugal a Preliminary View - Vítor Neves
1996-06	Housing, Mortgage Finance and the British Economy - Kenneth Gibb & Nile Istephan
1996-05	The Social Policy of The European Community, Reporting Information to Employees, a U.K. perspective: Historical Analysis and Prognosis - Ken Shackleton
1996-04	O Teorema da Equivalência Ricardiana: aplicação à economia portuguesa - Carlos Fonseca Marinheiro
1996-03	O Teorema da Equivalência Ricardiana: discussão teórica - Carlos Fonseca Marinheiro
1996-02	As taxas de juro no MMI e a Restrição das Reservas Obrigatórias dos Bancos - Fátima Assunção Sol e José Alberto Soares da Fonseca
1996-01	Uma Análise de Curto Prazo do Consumo, do Produto e dos Salários - João Sousa Andrade