Iterated failure rate monotonicity and ordering relations within Gamma and Weibull distributions*

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Abstract

Stochastic ordering of random variables may be defined by the relative convexity of the tail functions. This has been extended to higher order stochastic orderings, by iteratively reassigning tail-weights. The actual verification of stochastic orderings is not simple, as this depends on inverting distribution functions for which there may be no explicit expression. The iterative definition of distributions, of course, contributes to make that verification even harder. We have a look at the stochastic ordering, introducing a method that allows for explicit usage, applying it to the Gamma and Weibull distributions, giving a complete description of the order relations within each of these families.

1 Introduction

Ageing and ordering notions between random variables have long attracted the interest of a wide community. These notions raise intricate theoretical problems and have been widely used in applications in reliability, actuarial science, psychology, or economics (see, for example, Chandra and Roy [7], Nanda, Singh, Misra and Paul [20], Franco, Ruiz and Ruiz [13], Chechile [8], Belzunce, Candel and Ruiz [3, 4], Colombo and Labrecciosa [9], Veres-Ferrer and Pavia [32]). A connection with risk function properties may be found through utility functions, which may be interpreted as distribution functions. Ageing notions are usually defined in terms of the monotonicity of the survival or of the failure rate functions, while orderings between random variables, or to be more precise, their distributions, use relationships between these type of functions. The simplest ordering notions are based on direct comparisons between the survival or the failure rate functions. More interesting ordering relations, generally known as convex orderings, compare the decreasing rate of the tail functions through the relative convexity between the inverse tail functions. These convex orderings have been introduced by Hardy, Littlewood and Pólya [15], with some more recent results may be found in Palmer [23] or Rajba [28]. This means that the actual verification of these relations for given families of distributions is, in general, not obvious if the characterization of the distribution function are not simple, as is the case, for example, of the Gamma distributions. A classical early reference on ageing and some ordering problems for random variables and also on applications to reliability is the book by Barlow and Proschan [2]. More recent references on ageing and ordering notions, describing a nice account of properties and relations, including convex order notions, may be found in Mashall and Olkin [19] or Shaked and Shanthikumar [31].

A classification of families of distributions with respect to ageing notions was studied in Deshpande, Kochar and Singh [10] or Deshpande, Singh, Bagai and Jain [11]. Some of the classifications were based on higher order stochastic dominance, defined through relations between distributions constructed by iteratively reassigning their tail-weights as measures for the tails, as described in Definitions 6 and 14 below, and looking at the monotonicity of the failure rate functions after iteration. Those iterated relations were also studied by Averous and Meste [1], giving an almost complete picture of the ageing notions classification. The main focus being in establishing an hierarchy among the ageing notions rather than being very much concerned with the calculatory

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aspects. Naturally, the computational side of the problem becomes increasingly difficult, as a result of iterating the distribution functions. This means that, in general, simple questions as deciding whether a given distribution satisfies the appropriate monotonicity property is not simple. The monotonicity of failure rate functions after iteration has been considered in the literature, although under different names. Indeed, as mentioned in Nanda, Hazra, Al-Mutairi and Ghitany [21], the monotonicity of the twice iterated failure rate corresponds to the monotonicity in reversed direction of the mean residual life, studied by Bryson and Siddiqui [6], or, considering three iteration steps to the monotonicity, again in reversed direction, of the variance of the residual life, studied by Launer [18]. Hence, failure rate monotonicity after further iteration steps may be interpreted as the monotonicity of higher order moments of the residual life. Although starting from a different initial definition, the iterated distributions have also been used to compare different portfolios, as follows from the closed form representation we derive below. Indeed, the iterated distributions are, up to a constant factor, used to define Zolotarev ideal metrics (see, for example, Rachev and Rüschendorf [26] or the book by Rachev, Stoyanov and Fabozzi [25]). Ideal metrics appear as a natural way to measure the distance between different portfolios and depend on higher order truncated moments of the distributions. On other directions, Hanin and Rachev [14] applied ideal metrics to mass transshipment problems, Rachev and Rüschendorf [27] used the ideal metrics for the estimation of the Berry-Esséen bounds in compound Poisson models, or, more recently, Boutsikas and Vaggelatou [5] used this approach to characterize normal approximations under suitable dependence notions.

As what concerns ordering notions based in the tail-weight iterated distributions, a first classification study is found in Fagiuoli and Pellerey [12]. Again, the main concern is in establishing different relations between the several ordering notions, essentially with no explicit examples. The same problem, considering some new ordering notions was recently studied by Nanda et al. [21]. Once more, the main interest is in studying relations between the different orderings defined, with no examples. It is interesting to note that, although there is a vast literature on ordering (and ageing) notions, the actual verification of these relations is surprisingly difficult, even for the important and popular Gamma distributions (see, for example, Khaledi, Farsinezhadb and Kochar [16] or Kochar and Xu [17] for recent results on ordering relations within the Gamma family of distributions). Iterated order relations have been used by Sengupta and Deshpande [30] to obtain estimates for the difference between moments of random variables. Some results in this direction may also be found in [27].

We will look at ageing and convex ordering notions having in mind the purpose of introducing an actual computationally usable methodology to decide about the iterated failure rate monotonicity and ordering relations. The paper is organized as follows: Section 2 introduces the iterated distributions, gives a closed representation and defines the ageing notions, Section 3 establishes the failure rate ageing for the Weibull and Gamma distributions, Section 4 defines the convex order relation the paper is studying and proves characterizations in terms that are computationally exploitable. Finally, on Section 5 we use the previous results to give an almost complete classification of the order relations within the Gamma and Weibull families of distributions.

2 Definitions and basic representations

Let X be a nonnegative random variable with density function f_X, distribution function F_X, and tail function, or, as many authors call it, survival function, $\overline{F}_X = 1 - F_X$. We will be interested in ageing properties depending on iterated tail-weights for the distributions, as introduced by Averous and Meste [1] and initially studied by Fagiuoli and Pellerey [12].

Definition 1 *For each $x \geq 0$, define*

$$\overline{T}_{X,0}(x) = f_X(x) \quad \text{and} \quad \widetilde{\mu}_{X,0} = \int_0^\infty \overline{T}_{X,0}(t)\, dt = 1. \tag{1}$$

For each $s \geq 1$, define the $s-$iterated distribution $T_{X,s}$ by their tails $\overline{T}_{X,s} = 1 - T_{X,s}$ as follows:

$$\overline{T}_{X,s}(x) = \frac{1}{\widetilde{\mu}_{X,s-1}} \int_x^\infty \overline{T}_{X,s-1}(t)\, dt, \quad \text{where} \quad \widetilde{\mu}_{X,s-1} = \int_0^\infty \overline{T}_{X,s-1}(t)\, dt, \tag{2}$$

assuming this integral is finite. Moreover, we extend the domain of definition of each $\overline{T}_{X,s}$ defining $\overline{T}_{X,s}(x) = 1$, for $x < 0$.

We will be using these iterated distributions to establish ageing properties of distributions and ageing relations between different distributions within the same family. Our main concern is to introduce and use a method that actually allows the derivation of properties for specific families of distributions. For this purpose, we will be exploring a closed representation for the iterated distributions.

Lemma 2 *The tails $\overline{T}_{X,s}$ may be represented as*

$$\overline{T}_{X,s}(x) = \frac{1}{\prod_{j=1}^{s-1} \widetilde{\mu}_{X,j}} \int_x^\infty \frac{(t-x)^{s-1}}{(s-1)!} f_X(t)\, dt. \tag{3}$$

Proof. Successively replacing each $\overline{T}_{X,j}$, $j = s-1, \ldots, 1$, by its integral representation and reversing the integration order, we have

$$\begin{aligned}
\overline{T}_{X,s}(x) &= \frac{1}{\widetilde{\mu}_{X,s-1}} \int_x^\infty \frac{1}{\widetilde{\mu}_{X,s-2}} \int_t^\infty \overline{T}_{X,s-2}(u)\, du\, dt \\
&= \frac{1}{\widetilde{\mu}_{X,s-1}\widetilde{\mu}_{X,s-2}} \int_x^\infty \int_x^u \overline{T}_{X,s-2}(u)\, dt\, du \\
&= \frac{1}{\widetilde{\mu}_{X,s-1}\widetilde{\mu}_{X,s-2}} \int_x^\infty (u-x)\overline{T}_{X,s-2}(u)\, du \\
&= \frac{1}{\widetilde{\mu}_{X,s-1}\widetilde{\mu}_{X,s-2}} \int_x^\infty (u-x) \frac{1}{\widetilde{\mu}_{X,s-3}} \int_u^\infty \overline{T}_{X,s-3}(t)\, dt\, du \\
&= \frac{1}{\widetilde{\mu}_{X,s-1}\widetilde{\mu}_{X,s-2}\widetilde{\mu}_{X,s-3}} \int_x^\infty \frac{(t-x)^2}{2} \overline{T}_{X,s-3}(t)\, dt \\
&= \cdots = \frac{1}{\prod\limits_{j=1}^{k} \widetilde{\mu}_{X,s-j}} \int_x^\infty \frac{(t-x)^{k-1}}{(k-1)!} \overline{T}_{X,s-k}(t)\, dt.
\end{aligned}$$

So, finally, taking $k = s$, we obtain

$$\overline{T}_{X,s}(x) = \frac{1}{\prod\limits_{j=1}^{s} \widetilde{\mu}_{X,s-j}} \int_x^\infty \frac{(t-x)^{s-1}}{(s-1)!} f_X(t)\, dt.$$

To conclude the proof, reverse the indexing order on the product of the $\widetilde{\mu}_{X,j}$. ■

Remark 3 *From (3), if we choose $x = 0$ and take into account that $\overline{T}_{X,s}(0) = 1$, it follows that*

$$\mathbb{E}X^{s-1} = (s-1)! \prod_{j=0}^{s-1} \widetilde{\mu}_{X,j}.$$

Replacing this expression in (3), another representation of $\overline{T}_{X,s}$ follows:

$$\overline{T}_{X,s}(x) = \frac{1}{\mathbb{E}X^{s-1}} \int_x^\infty f_X(t)(t-x)^{s-1}\, dt. \tag{4}$$

Moreover, it also follows an explicit expression for moments of the iterated distributions:

$$\widetilde{\mu}_{X,s-1} = \frac{1}{s-1} \frac{\mathbb{E}X^{s-1}}{\mathbb{E}X^{s-2}}.$$

Remark 4 *The representation (4) means that $\overline{T}_{X,s} = \frac{1}{\mathbb{E}X^{s-1}} \mathbb{E}(X-x)_+^{s-1}$. Given random variables X and Y, representing portfolios, the stop-loss distance or Zolotarev ideal metric of order s between the portfolios (see, for example, Rachev and Rüschendorf [26] or Rachev et al. [25]) may be written as $\mathbb{E}X^{s-1}\overline{T}_{X,s} - \mathbb{E}Y^{s-1}\overline{T}_{Y,s}$.*

We now discuss some definitions of ageing. One of the most simple and common ageing notions is based on the failure rate function of a distribution $\frac{f_X(x)}{1-F_X(x)} = \frac{\overline{T}_{X,0}(x)}{\overline{T}_{X,1}(x)}$. Even before getting into comparisons between probability distributions, studied later in this paper, the monotonicity of the failure rate function is a relevant property, satisfied by many common distributions. The direct verification of this monotonicity may not be a simple task, as for many distributions the tail does not have an explicit closed representation or, at least, not a manageable one. As we have defined iterated distributions, it becomes now natural to proceed likewise with respect to the failure rate functions.

Definition 5 *For each $s \geq 1$ and $x \geq 0$, define the $s-$iterated failure rate function as*

$$r_{X,s}(x) = \frac{\overline{T}_{X,s-1}(x)}{\int_x^\infty \overline{T}_{X,s-1}(t)\,dt} = \frac{\overline{T}_{X,s-1}(x)}{\widetilde{\mu}_{X,s-1}\overline{T}_{X,s}(x)}.$$

It is obvious that for $s = 1$, we find the failure rate of X:

$$r_{X,1}(x) = \frac{\overline{T}_{X,0}(x)}{\widetilde{\mu}_0\overline{T}_{X,1}(x)} = \frac{f_X(x)}{1 - F_X(x)} = \frac{f_X(x)}{\overline{F}_X(x)}.$$

Thus the monotonicity of the failure rate is expressed as the monotonicity of $r_{X,1}$. We may extend this monotonicity notion by considering the $s-$iterared distribution, as done in Averous and Meste [1] and Fagiuoli and Pellerey [12], among many other authors.

Definition 6 *For $s = 1, 2, \ldots$, the nonnegative random variable X is said to be*

1. *$s-$IFR (resp., $s-$DFR) if $r_{X,s}$ is increasing (resp., decreasing) for $x \geq 0$.*
2. *$s-$IFRA if $\frac{1}{x}\int_0^x r_{X,s}(t)\,dt$ is increasing for $x > 0$.*
3. *$s-$NBU if $\overline{T}_{X,s}(x+t) \leq \overline{T}_{X,s}(x)\overline{T}_{X,s}(t)$, for all $x, t \geq 0$.*
4. *$s-$NBUFR if $r_{X,s}(0) \leq r_{X,s}(x)$, for all $x \geq 0$.*
5. *$s-$NBAFR if $r_{X,s}(0) \leq \frac{1}{x}\int_0^x r_{X,s}(t)\,dt$, for all $x > 0$.*

Throughout this paper, we will be interested mainly in the $s-$IFR notion. But, as proved by Fagiuoli and Pellerey [12], this is the strongest notion. The following lemma states the relevant part, for the purposes of the present paper, of the relations between the above notions proved by Fagiuoli and Pellerey [12] (see their Figure 2 for an easily readable account of the relations proved).

Lemma 7 *Let X be a nonnegative random variable with finite $s - 1$ moment. Then, for each integer $s \geq 1$, the following implications hold: X is $s-$IFR $\Rightarrow$ X is $s-$IFRA $\Rightarrow$ X is $s-$NBU $\Rightarrow$ X is $s-$NBUFR $\Rightarrow$ X is $s-$NBAFR.*

3 Iterated failure rate monotonicity

The iterated failure rate property of a distribution turns out not to add much to the ageing notion. Indeed, it follows from the results in Fagiuoli and Pellerey [12], that, if the appropriate moments exist, higher order monotonicity of the failure rate are inherited from the corresponding lower order properties, as described next.

Lemma 8 *Let X be a nonnegative random variable with finite moment of order $s \geq 1$. The following implications are true:*

a) If X is $s-$IFR, then X is $(s+1)-$IFR.

b) If X is $s-$DFR, then X is $(s+1)-$DFR.

Proof. This is an immediate consequence of Theorems 3.4 and 4.3 in Fagiuoli and Pellerey [12]. ■

This result is included in Theorem 2 in Navarro and Hernandez [22]. It implies that, for most distributions it is enough to verify the 1−IFR or the 1−DFR property. However, we may find examples of distributions for which lower order monotonicity does not hold, but this property is true after a few iterations.

Example 9 *Consider a random variable X with density $f_X(x) = c' \frac{\log(x+c)}{(x+c)^3}$, for $x > 0$, where c' is a suitable normalizing constant. This is a fattened tail Pareto type distribution. Simple integration shows that,*

$$\overline{T}_{X,1}(x) = c' \frac{2\log(x+c)+1}{4(x+c)^2} \quad \text{and} \quad \overline{T}_{X,2}(x) = \frac{c}{2\log c + 3} \frac{2\log(x+c)+3}{x+c},$$

hence

$$r_{X,1} = \frac{4\log(x+c)}{(x+c)(2\log(x+c)+1)} \quad \text{and} \quad r_{X,2} = \frac{2\log(x+c)+1}{(x+c)(2\log(x+c)+3)}.$$

It is now easily verified that if $c \in (\exp(-1+\sqrt{20}/4), e^{1/2}) \approx (1.12528, 1.64872)$, the 1−iterated failure rate $r_{X,1}$ is not monotone, while the 2−iterated failure rate $r_{X,2}$ is decreasing.

The density considered above has finite expectation, but the second order moment does not exist. Therefore, the 3−iterated distribution is no longer definable. To obtain examples of distributions with higher order finite moments for which we have the same behaviour for the monotonicity as above for $r_{X,1}$ and $r_{X,2}$, consider densities of the form $c' \frac{\log(x+c)}{(x+c)^\alpha}$, for some large enough $\alpha > 0$. It is possible to check the existence of an interval for the choice of the parameter c where the 1−iterated failure rate is not monotonous and the 2−iterated failure rate is decreasing, although the explicit characterizations for this interval becomes somewhat cumbersome. Naturally, the same approach is possible with higher moments to obtain a density such that the 2−iterated failure is not monotonous and the 3−iterated failure rates is decreasing, and so on to construct distributions that only show failure rate monotonicity properties after some iteration steps.

We will now use the property stated in Lemma 8 to describe the failure rate monotonicity of the Weibull and the Gamma families of distributions. We prove here the complete result for the iterated failure rate monotonicity of the Weibull and of the Gamma distributions.

Theorem 10 *Let X be a nonnegative random variable with Weibull distribution with shape parameter α and scale parameter θ, and $s \geq 1$ an integer. If $\alpha \geq 1$ (resp., $\alpha < 1$), then X is s−IFR (resp., s−DFR).*

Proof. Taking into account Lemma 8, it is enough to consider the case $s = 1$. Using the expression for the distribution function of X, it follows that the quotient $r_{X,1}(x) = \frac{f_X(x)}{\overline{F}_X(x)} = \frac{\alpha}{\theta}\left(\frac{x}{\theta}\right)^{\alpha-1}$, which is increasing if $\alpha \geq 1$ and decreasing otherwise. ■

We now handle the Gamma distributions. For this family of distributions, we cannot compute explicitly the failure rate function, as in general, the distribution function does not have a closed form representation. So, we need a work around to prove the monotonicity. Let us start by stating without proof, a simple but useful characterization of monotonicity.

Lemma 11 *A function $g : \mathbb{R} \longrightarrow \mathbb{R}$ is increasing (resp., decreasing) if and only if for every $a \in \mathbb{R}$, $g(x) - a$ changes sign at most once when x traverses from $-\infty$ to $+\infty$, and if the change occurs, it is in the order "$-,+$" (resp., "$+,-$").*

Theorem 12 *Let X be a nonnegative random variable with distribution $\Gamma(\alpha, \theta)$, and $s \geq 1$ an integer. If $\alpha \geq 1$, then X is s−IFR. If $\alpha < 1$, then X is s−DFR.*

Proof. Again, from Lemma 8, it is enough to prove that X is either 1−IFR or 1−DFR, that is, to prove the increasingness or decreasingness of the quotient $r_{X,1}(x) = \frac{f_X(x)}{\overline{F}_X(x)}$. Since, in general, there is no explicit closed form for $\overline{F}_X(x)$, we will prove the monotonicity using Lemma 11. As

f_X and $\overline{F}_X$ are nonnegative, it is enough to take, while applying Lemma 11, the constant $a > 0$. So, for every given $a > 0$, we shall study the sign variation of $\frac{f_X(x)}{\overline{F}_X(x)} - a$. Remark that the sign of this difference coincides, for every $x \geq 0$, with the sign of $H(x) = f_X(x) - a\overline{F}_X(x)$, so it is enough to study the sign variation of H. It is obvious that $H(0) = -a < 0$ and, if $\alpha \geq 1$ we have $\lim_{x\to+\infty} H(x) = 0$. Now, differentiating, we find that

$$H'(x) = f'_X(x) + af_X(x) = \frac{x^{\alpha-2}e^{-x/\theta}}{\theta^{\alpha+1}\Gamma(\alpha)}((a\theta-1)x - \theta(1-\alpha)),$$

so the sign of H' is determined by the sign of the straight line $\ell(x) = (a\theta-1)x - \theta(1-\alpha)$. Obviously $\ell(0) = -\theta(1-\alpha)$. Keeping in mind that we are assuming that $\alpha \geq 1$, it follows that $\ell(0) > 0$, thus the sign variation of $\ell(x)$ in $[0, +\infty)$ is "+", if $(a\theta - 1) > 0$, or "$+, -$", if $(a\theta - 1) < 0$. In the first case, where $\ell(x) > 0$, for all $x > 0$, the function H is always increasing so, given its value at 0 and at infinity, the sign of H is "$-$". In the case where the sign variation of ℓ is "$+, -$", again taking into account the behaviour of H at the origin and at infinity, implies that its sign variation is "$-, +$". The case $\alpha < 1$ is analysed analogously by taking into account that $\lim_{x\to 0} H(x) = +\infty$. ■

The $s-$IFR-ness of the Gaussian distributions is proved in an analogous way. We state the result without proof.

Theorem 13 *Gaussian distributions are $s-$IFR.*

4 Iterated failure rate ordering

We now compare different distributions with respect to their iterated failure rate monotonicity rates. In the sequel, let $\mathcal{F}$ denote the family of distributions functions such that $F(0) = 0$ and the corresponding probability distribution has support contained in $[0, +\infty)$. In this section, we will define an iterated failure rate order and prove a general criterium. We start by defining the ordering, following Nanda et al. [21].

Definition 14 *Let X and Y be random variables with distribution functions $F_X, F_Y \in \mathcal{F}$ and $s \geq 1$ an integer. The random variable X (or its distribution F_X) is said to be less $s-$IFR than Y (or its distribution F_Y), and we write $X \leq_{s-IFR} Y$, or equivalently, $F_X \leq_{s-IFR} F_Y$, if $c_s(x) = \overline{T}_{Y,s}^{-1}(\overline{T}_{X,s}(x))$ is convex.*
Moreover, two nonnegative random variables X and Y, or two distribution functions F_X, $F_Y \in \mathcal{F}$, are said to be $s-$IFR equivalent, denoted by $X \sim_{s-IFR} Y$ or $F_X \sim_{s-IFR} F_Y$, if there exists a constant $k > 0$ such that $F_X(x) = F_Y(kx)$, for all $x \geq 0$.

The $s-$IFR relation between random variables, or their distributions to be more precise, is a variant of relative convexity between real functions f_1 and f_2, as defined by Hardy et al. [15], Pečarić, Proschan and Tong [24] or Roberts and Varberg [29]. However, these authors define the relative convexity of f_1 with respect to f_2 requiring the convexity of $f_1(f_2^{-1}(x))$, that is, using the inverse functions in reversed order when compared to Definition 14. For some more recent results on the characterization of relative convexity, we may refer the reader to Palmer [23] or Rajba [28]. These authors give some equivalent characterizations of relative convexity, but they all depend on the functions that are compared. So, if we do not have closed and manageable representations for these functions, as is the case for the distributions we will be analysing below, the effective calculation difficulty remains.

It is possible to define several other ordering relations corresponding to the different ageing notions referred in Definition 6, as was done by Nanda et al. [21]. We will be only interested in the $s-$IFR ordering, so we do not quote here those other ordering notions. Results and characterizations similar to the ones we derive below may also be proved for these alternative stochastic order relations. Moreover, as happens for the ageing notions, the $s-$IFR ordering is the strongest of those order relations, as proved in Nanda et al. [21]. We have been referring to the $s-$IFR as an ordering but, of course, one has to verify that this is really the case. This has been proved in Nanda et al. [21].

Lemma 15 (Theorem 2.1 in Nanda et al. [21]) *The relationship $F_X \leq_{s-IFR} F_Y$ defines an order relation on the equivalence classes with respect to $\sim_{s-IFR}$, of $\mathcal{F}$.*

Example 16 *This order relation is indeed only partial, as shown by the following example. Consider X with inverse Gamma distribution with shape parameter $\alpha = 1$ and scale parameter $\beta > 0$, Y with exponential distribution with scale parameter λ, and consider $s = 1$. Then, we have:*

$$f_X(x) = \frac{\beta}{x^2} e^{-\beta/x}, \qquad \overline{T}_{X,1}(x) = \overline{F}_X(x) = 1 - e^{-\beta/x}, \quad \text{and} \quad \overline{T}_{Y,s}(x) = e^{-x/\lambda},$$

so,

$$c_1(x) = -\lambda \log \overline{F}_X(x) = -\lambda \log(1 - e^{-\beta/x}),$$

is neither convex nor concave, thus X and Y are not comparable with respect to $1-$IFR.

From Definition 14 and Lemma 15, it follows immediately that the multiplying random variables by positive constants will not affect the $s-$IFR ordering relation.

Corollary 17 *Let X and Y be nonnegative random variables with distributions $F_X, F_Y \in \mathcal{F}$, $s \geq 1$ an integer, and $\alpha_1, \alpha_2 > 0$. $X \leq_{s-IFR} Y$ if and only if $\alpha_1 X \leq_{s-IFR} \alpha_2 Y$.*

The previous result will be useful to compare parametric distributions where there exists a scale parameter, as it follows that we may assume this parameter to be equal to 1.

The exponential distribution plays an important role when dealing with ageing notions. As already proved by Nanda et al. [21], the $s-$IFR comparability with the exponential is equivalent to the failure rate monotonicity.

Theorem 18 (Theorem 2.2 in Nanda et al. [21]) *Let X be a random variable with distribution function $F_X \in \mathcal{F}$ and Y with exponential distribution with scale parameter λ. Then $X \leq_{s-IFR} Y$ (resp., $Y \leq_{s-IFR} X$) if and only if X is $s-$IFR (resp., X is $s-$DFR).*

Proof. We have $\overline{T}_{Y,s} = e^{-x/\lambda}$, thus $F_X \leq_{s-IFR} F_Y$ is equivalent to $\log(\overline{T}_{X,s}(x))$ being concave, which is equivalent to requiring that $r_{X,s}(x)$ is increasing, that is, X is $s-$IFR. ■

Remark 19 *As what regards the comparison of two distribution with respect to the $s-$IFR order an inheritance relation suggested by the Lemma 8 could be expected. We will not prove any such inheritance relationship. Moreover, the previous result allows to exhibit examples of distributions for which no $1-$IFR order exists, but a $2-$IFR order does exist. Indeed, consider, as in Example 9, the random variable X with density $f_X(x) = c' \frac{\log(x+c)}{(x+c)^3}$, for $x > 0$. Taking into account Theorem 18, for $c \in (1.12528, 1.64872)$, it follows that X is not comparable with any exponential distribution with respect to the $1-$IFR ordering. On the other hand, the decreasingness of $r_{X,2}$, proved in Example 9, also shows that the exponential distribution is less $2-$IFR than X. More generally, this ordering remains true in the $s-$IFR sense, for $s \geq 2$.*

Note that the characterization described in Theorem 18 provides an alternative way to our statements about the $s-$IFR-ness of the Weibull and Gamma distributions (Theorems 10 and 12 above). Of course, to use Theorem 18, we still need an effective way to compare distributions with respect to $s-$IFR order relation, and this may not be a simple task. Indeed, the direct verification of the convexity of c_s, stated in Definition 14, is in general difficult to perform, as we cannot find explicit closed representations of the distributions functions involved in the definition of c_s, thus we cannot invert $\overline{T}_{Y,s}$. One could try to use the characterization of the derivative of the inverse function for this purpose. This is exactly what was done in Proposition 2.2 in Nanda et al. [21] to obtain alternative characterizations for the $s-$IFR ordering. But those alternatives are not really effective for actual computation purposes, as they all depend on monotonicity relations of transformations of the iterated distribution functions and their inverses. Thus, in all cases where no explicit closed representations is available, as for the Gamma family, we still have no effective way to conclude about the order relation. As already commented above, the characterizations proved by Palmer [23] or Rajba [28] do not help on this matter.

We shall start by proving an alternative characterization for the convexity of a continuous real function in terms of crossings of their graphical representations with straight lines.

Theorem 20 *Let f be a continuous function. The function f is convex if and only if for every real numbers a and b, $f(x) - (ax + b)$ changes sign at most twice when x traverses from $-\infty$ to $+\infty$, and if the change of sign occurs twice, it is in the order "$+, -, +$".*

Proof. Assume that $f(x) - (ax + b)$ changes sign more than twice or in the order "$-, +, -$". In both cases there exists an interval where the sign change sequence is in the order "$-, +, -$". But this means that the function f is not convex as, after getting above a straight line it crosses again under the same straight line.

Assume now that f is not convex, then there exists an interval $I = [x_0, x_1]$ such that $f(x) > \frac{f(x_1)-f(x_0)}{x_1-x_0}(x - x_0) + f(x_0)$, for all $x \in (x_0, x_1)$, that is, the graph of f is, for $x \in I$, above the line Δ defined by $(x_0, f(x_0))$ and $(x_1, f(x_1))$. Let Δ_ε be the line obtained by shifting upwards Δ by ε, described by the equation $y = \frac{f(x_1)-f(x_0)}{x_1-x_0}(x - x_0) + f(x_0) + \varepsilon$. It is obvious that, at least for ε small enough, the sign variation of $f(x) - \left(\frac{f(x_1)-f(x_0)}{x_1-x_0}(x - x_0) + f(x_0) + \varepsilon\right)$, for $x \in I$, is at least in the order "$-, +, -$". ∎

To complement the previous result, the following characterization of the crossing of two graphical representations will be useful.

Lemma 21 (Marshal and Olkin [19], pp. 699–700) *Let f and g two real-valued functions, and ζ be a strictly increasing (resp., decreasing) and continuous function defined on the range of f and g. For any real number $c > 0$, the functions $f(x) - cg(x)$ and $\zeta(f(x)) - \zeta(cg(x))$ have the same (resp., reverse) sign variation order as x traverses from $-\infty$ to $+\infty$.*

The previous results provide an immediate and simple alternative characterization of $s-$IFR order relation.

Theorem 22 *Let X and Y be random variables with distribution functions $F_X, F_Y \in \mathcal{F}$. $X <_{s-IFR} Y$ if and only if for any real numbers a and b, $\overline{T}_{Y,s}(x) - \overline{T}_{X,s}(ax + b)$ changes sign at most twice, and if the change of signs occurs twice, it is in the order "$+, -, +$", as x traverses from 0 to $+\infty$.*

Remark 23 *We have reduced the variation of x to traversing from 0 to $+\infty$ because all the functions $\overline{T}_{X,s}$ and $\overline{T}_{Y,s}$ are equal to 1 for $x < 0$.*

Definition 24 *Given random variables X and Y, we denote $V_s(x) = \overline{T}_{Y,s}(x) - \overline{T}_{X,s}(ax + b)$.*

It is obvious from the definition of the iterated tails that V_s is differentiable.

Remark 25 *Taking into account that, being the tails of distributions, $\overline{T}_{X,s}$ and $\overline{T}_{Y,s}$ are decreasing it is enough to consider, when applying Theorem 22, the constant $a > 0$. Indeed, we have $V_s(0) = 1 - \overline{T}_{X,s}(b)$, and if $a < 0$,*

$$V_s'(x) = -\frac{1}{\widetilde{\mu}_{Y,s-1}}\overline{T}_{Y,s-1}(x) + \frac{a}{\widetilde{\mu}_{X,s-1}}\overline{T}_{X,s-1}(ax + b) \leq 0.$$

Now, it is obvious that for $a < 0$, we have $\lim_{x\to+\infty} V_s(x) = -1$. Thus, the sign variation of V will be "$+, -$", if $b > 0$, and "$-$" if $b \leq 0$. That is, in both cases we meet the convexity condition described in Theorem 20.

Finally, we prove a simple result describing the sign variation after performing integration. This will be convenient for the later discussion.

Lemma 26 *Let f and g be two real-valued functions defined on $[0, \infty)$ such that,*

$$g(x) = \int_x^\infty f(t)\,dt.$$

Assume that, as x traverses from 0 to $+\infty$, $f(x)$ changes sign in one of the following orders "$-, +$" or "$+, -$" or "$+, -, +$" or "$-, +, -, +$". Then $g(x)$, as x traverses from 0 to ∞, has sign variation equal to every possible final part of the sign variation of $f(x)$.

Proof. The proof follows from a simple argument using that $g'(x) = -f(x)$, and separating into the four possible sign variations considered. ■

We may now prove a general criterium to compare with respect to the $s-$IFR order two distribution functions.

Theorem 27 *Let X and Y be random variables with absolutely continuous distributions with densities f_X and f_Y, and distribution functions $F_X, F_Y \in \mathcal{F}$, respectively. If, for some positive integer $k \leq s$, and every $a > 0$ and $b \in \mathbb{R}$, the function*

$$H_k(x) = \frac{1}{\prod_{j=1}^{k} \widetilde{\mu}_{Y,s-j}} \overline{T}_{Y,s-k}(x) - \frac{a^k}{\prod_{j=1}^{k} \widetilde{\mu}_{X,s-j}} \overline{T}_{X,s-k}(ax+b) \tag{5}$$

changes sign at most twice, and if the change of signs occurs twice, it is in the order "$+,-,+$", as x traverses from 0 to $+\infty$, then $F_X \leq_{s-IFR} F_Y$.

Proof. Remember the integral representation for $V_s(x) = \overline{T}_{Y,s}(x) - \overline{T}_{X,s}(ax+b)$ obtained in the intermediate steps of the proof of Lemma 2:

$$\begin{aligned} V_s(x) &= \frac{1}{\prod_{j=1}^{k} \widetilde{\mu}_{Y,s-j}} \int_x^{\infty} \frac{(t-x)^{k-1}}{(k-1)!} \overline{T}_{Y,s-k}(t)\, dt \\ &\qquad - \frac{1}{\prod_{j=1}^{k} \widetilde{\mu}_{X,s-j}} \int_{ax+b}^{\infty} \frac{(t-(ax+b))^{k-1}}{(k-1)!} \overline{T}_{X,s-k}(t)\, dt \\ &= \int_x^{\infty} \frac{(t-x)^{k-1}}{(k-1)!} H_k(t)\, dt, \end{aligned}$$

after an appropriate change of variable in the second integral. Now, using Theorem 22 and Lemma 26 the proof is concluded. ■

Remark 28 *As mentioned before, in general, the explicit forms of $\overline{T}_{X,s}$ and $\overline{T}_{Y,s}$ are difficult to obtain. So, in most of the cases we will be interested in applying Theorem 27 choosing $k = s$, thus using the density functions to define H_s, or $k = s-1$, using the distribution functions to define H_{s-1}, if these are available.*

Following the previous remark, we have a closer look at H_s and H_{s-1}, and the control of their sign variation. Taking into account the representation (4) for the iterated tails, we have

$$H_s(x) = \frac{1}{\mathbb{E}Y^{s-1}} f_Y(x) - \frac{a^s}{\mathbb{E}X^{s-1}} f_X(ax+b)$$

and,

$$H_{s-1}(x) = \frac{1}{\mathbb{E}Y^{s-1}} \overline{F}_Y(x) - \frac{a^{s-1}}{\mathbb{E}X^{s-1}} \overline{F}_X(ax+b).$$

In most cases, the direct analysis of the sign variation of H_s is, to say the least, difficult, even for relatively simple density functions as, for example, the Gamma densities. An alternative approach to the control of this sign variation is to apply Lemma 21, choosing an appropriate ζ transformation. For the family of distributions we will be considering in the sequel, we shall take $\zeta(x) = \log x$.

Corollary 29 *Let X and Y be random variables with absolutely continuous distributions with densities f_X and f_Y and distribution functions $F_X, F_Y \in \mathcal{F}$, respectively. If, for every constants $a > 0$ and $b \in \mathbb{R}$, either of the functions,*

$$P_s(x) = \log f_Y(x) - \log f_X(ax+b) + \log \frac{\mathbb{E}X^{s-1}}{a^s \mathbb{E}Y^{s-1}},$$

or,

$$P_{s-1}(x) = \log \overline{F}_Y(x) - \log \overline{F}_X(ax+b) + \log \frac{\mathbb{E}X^{s-1}}{a^{s-1} \mathbb{E}Y^{s-1}},$$

changes sign at most twice when x traverses from 0 to $+\infty$, and if the change of sign occurs twice it is in the order "$+,-,+$", then $F_X \leq_{s-IFR} F_Y$.

5 Some applications

In this section, we will be applying the general characterizations derived before to establish the $s-$IFR ordering among some families of distributions, and analyse an example highlighting the influence of the various parameters appearing in the expression of P_s and P_{s-1}.

5.1 Comparing two Gamma distributions

As argued after Corollary 17, it is enough to compare Gamma distributions both with the same scale parameter $\theta = 1$. We will be using Corollary 29 with respect to P_s, assuming X has $\Gamma(\alpha', 1)$ distribution and Y has $\Gamma(\alpha, 1)$ distribution. Thus, we need to analyse the sign variation in $[0, +\infty)$ of

$$P_s(x) = (\alpha - 1)\log x - (\alpha' - 1)\log(ax + b) - x + ax + b + \log\frac{\Gamma(\alpha')}{\Gamma(\alpha)} + \log\frac{\mathbb{E}X^{s-1}}{a^s\mathbb{E}Y^{s-1}}, \qquad (6)$$

where $a > 0$ and $b \in \mathbb{R}$. Note that $\lim_{x\to+\infty} P_s(x) = \infty \times \operatorname{sgn}(a-1)$. Differentiating the expression above, we have

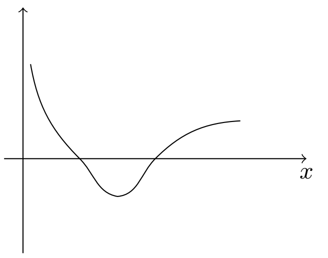
$$P_s'(x) = \frac{\alpha - 1}{x} - \frac{a(\alpha' - 1)}{ax + b} + a - 1 = \frac{a(a-1)x^2 + ((\alpha - \alpha')a + (a-1)b)x + (\alpha - 1)b}{x(ax + b)}. \qquad (7)$$

Let us denote the numerator in (7) by $N_s(x) = a(a-1)x^2 + ((\alpha - \alpha' -)a + (a-1)b)x + (\alpha - 1)b$. To analyse the sign variation of V_s, we need to separate between the cases when $b \geq 0$ and $b < 0$. Indeed, while for the first case, we need to consider x traversing from 0 to $+\infty$, for the later case, we will be only analysing the sign variation in the interval $(-\frac{b}{a}, +\infty)$ as, for $x \leq -\frac{b}{a}$, $V_s(x) = \overline{T}_{Y,s}(x) - 1 \leq 0$. Hence, for both cases, in the interval of interest, the sign of P_s' is determined by the sign of N_s.

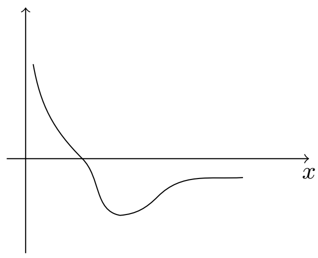
Proposition 30 *Let $\alpha' > \alpha > 1$ and $\theta_1, \theta_2 > 0$. The $\Gamma(\alpha', \theta_1)$ distribution is less $s-$IFR than the $\Gamma(\alpha, \theta_2)$ distribution.*

Proof. Taking into account Corollary 17, we may assume $\theta_1 = \theta_2 = 1$. Moreover, remember that, according to Remark 25, it is enough to take $a > 0$. Note further that $\lim_{x\to 0^+} P_s(x) = -\infty$.

Assume first that $b \geq 0$. The convexity of N_s is determined by the sign of $a - 1$, and $N_s(0) = (\alpha - 1)b \geq 0$, so the behaviour of P_s' may be as follows:



P_s' for $a > 1$



P_s' for $a < 1$

Case 1: $a > 1$. We have $\lim_{x\to+\infty} P_s(x) = +\infty$, thus the most sign varying situation corresponds to "$-, +, -, +$" implying, based on Lemma 26, that the sign variation of V_s might be "$-, +, -, +$" or "$+, -, +$" or "$-, +$" or "$+$". Now, as $V_s(0) = 1 - \overline{T}_{X,s}(b) \geq 0$, the only possible cases are "$+, -, +$" or "$+$".

Case 2: $a \leq 1$. In this case we have $\lim_{x\to+\infty} P_s(x) = -\infty$. The behaviour of P_s' when $a = 1$ is still described by the picture on the right, with P_s' approaching 0 as $x \longrightarrow +\infty$, instead of being strictly negative. Taking into account this behaviour of P_s' the monotonicity of P_s is, $\nearrow\searrow$, meaning that the most sign varying case for P_s is "$-, +, -$". Again, based on Lemma 26 and $V_s(0) \geq 0$, the only possible sign variation is "$+, -$".

Assume now that $b < 0$. Then, we have, for $x \leq -\frac{b}{a}$, $V_s(x) = \overline{T}_{Y,s}(x) - 1 \leq 0$, so it remains to describe the sign variation for $x > -\frac{b}{a}$, thus needing to locate $-\frac{b}{a}$ with respect to the roots of N_s. As $N_s(0) = (\alpha - 1)b < 0$, two situations may occur:

Case 3: $a > 1$. The sign variation of $N_s(x)$ in the interval $(-\frac{b}{a}, +\infty)$ is either "+" or "$-,+$". As $\lim_{x\to(-b/a)^+} P_s(x) = +\infty$ and $\lim_{x\to+\infty} P_s(x) = +\infty$, it follows that the sign variation of $N_s(x)$ is "$-,+$". Thus, the most sign varying possibility for P_s in the interval $(-\frac{b}{a}, +\infty)$ is "$+,-,+$". From Lemma 26, it follows that the sign variation for V_s in $(-\frac{b}{a}, +\infty)$ is one of the three possibilities: "$+,-,+$" or "$-,+$" or "+". As $V_s(-\frac{b}{a}) \leq 0$, it follows that the sign variation of V_s in $(0, +\infty)$ is "$-,+$".

Case 4: $a \leq 1$. The sign variation of $N_s(x)$ in the interval $(-\frac{b}{a}, +\infty)$ is either "$-$", or "$-,+,-$". Assume first that the sign variation of $N_s(x)$ is "$-,+,-$", which means that N_s has two positive roots and its maximum is reached for $x = -\frac{b}{2a} + \frac{\alpha-\alpha'}{2(1-a)} < -\frac{b}{a}$, therefore, the sign variation of N_s in the interval $(-\frac{b}{a}, +\infty)$ is "$+,-$" or "$-$". As $\lim_{x\to(-b/a)^+} P_s(x) = +\infty$, there is only one possible sign variation of N_s, which is "$-$". Hence, the sign variation of P_s is, at most, "$+,-$". Now, using the fact that $V_s(-\frac{b}{a}) \leq 0$, it follows the sign variation of V_s in $[0, +\infty)$ is "$-$". It remains to analyse the case where N_s is always negative, but the description of the sign variation of V_s follows in the same way.

So, finally, the possibilities for the sign variation of V_s are: either at most one sign change or, in case of two sign changes, these are "$+,-,+$". Hence, the conclusion follows taking into account Theorem 22. ∎

Proposition 31 *Let $\alpha' > 1 > \alpha > 0$ and $\theta_1, \theta_2 > 0$. The $\Gamma(\alpha', \theta_1)$ distribution is less $s-$IFR than the $\Gamma(\alpha, \theta_2)$ distribution.*

Proof. The result follows immediately using Theorem 18 and the transitivity of the $s-$IFR-order, by comparing both of them with the exponential distribution. ∎

5.2 Comparing two Weibull distributions

In the sequel, we shall denote by $W(\alpha, \theta)$ the Weibull distribution with shape parameter α and scale parameter θ. As for the Gamma family of distributions, it is enough to compare Weibull distributions both with scale parameter $\theta = 1$. Moreover, we will apply Corollary 29 now with respect to P_{s-1}, as the tail of a Weibull distribution has a simple closed form representation, assuming that X has distribution $W(\alpha', 1)$ and Y has distribution $W(\alpha, 1)$. So, we are interested in analysing the sign variation of

$$P_{s-1}(x) = -x^\alpha + (ax+b)^{\alpha'} + \log \frac{\mathbb{E}X^{s-1}}{a^{s-1}\mathbb{E}Y^{s-1}}, \tag{8}$$

where $a > 0$ and $b \in \mathbb{R}$. Differentiating this expression, we have

$$P'_{s-1} = a\alpha'(ax+b)^{\alpha'-1} - \alpha x^{\alpha-1}. \tag{9}$$

The direct control of the sign variation of P'_{s-1} is too difficult, so we will use again Lemma 21 with the choice $\zeta(x) = \log x$. This means that the sign variation of P'_{s-1} is the same as the sign variation of

$$\begin{aligned} Q_{s-1}(x) &= \log(a\alpha'(ax+b)^{\alpha'-1}) - \log(\alpha x^{\alpha-1}) \\ &= \log(a\alpha') - \log\alpha + (\alpha'-1)\log(ax+b) - (\alpha-1)\log x, \end{aligned} \tag{10}$$

whose derivative is

$$Q'_{s-1}(x) = \frac{a(\alpha'-1)}{ax+b} - \frac{\alpha-1}{x} = \frac{a(\alpha'-\alpha)x + (1-\alpha)b}{x(ax+b)}. \tag{11}$$

Proposition 32 *Let $\alpha' > \alpha > 1$ and $\theta_1, \theta_2 > 0$. The $W(\alpha', \theta_1)$ distribution is less $s-$IFR than the $W(\alpha, \theta_2)$ distribution.*

Proof. As before, without loss of generality, we may take $\theta_1 = \theta_2 = 1$ and use the representations (8)–(11). As usual, we need to separate the cases $b > 0$ and $b \le 0$, and remember that we need only to assume that $a > 0$.

Assume first that $b > 0$. It follows from (11) that the sign variation in the interval $(0, +\infty)$ for Q'_{s-1} is "$-, +$", hence the monotonicity of Q_{s-1}, in this same interval, is $\searrow\nearrow$. From (10), it follows that $\lim_{x\to 0^+} Q_{s-1}(x) = +\infty$ and $\lim_{x\to+\infty} Q_{s-1}(x) = +\infty$, so the most sign varying possibility for Q_{s-1}, which coincides with the sign variation of P'_{s-1}, is "$+, -, +$". It follows that the monotonicity of P_{s-1} is $\nearrow\searrow\nearrow$. Note that $\alpha' > \alpha$ implies $\lim_{x\to+\infty} P_{s-1}(x) = +\infty$. Hence, the most sign varying possibility in the interval $(0, +\infty)$ for P_{s-1} is "$-, +, -, +$". Taking now into account Lemma 26, the sign variation of V_s in the interval $(0, +\infty)$ may be "$-, +, -, +$" or "$+, -, +$" or "$-, +$" or "$+$", and remembering that $V_s(0) \ge 0$, the actual possible choices are "$+, -, +$" or "$+$".

Assume now that $b \le 0$. As explained before, we need only to describe the sign variation in $(-\frac{b}{a}, +\infty)$. Now from (11) it follows that, for $x > 0$, we have $Q'_{s-1}(x) > 0$, so Q_{s-1} is always increasing in $(0, +\infty)$. As $\alpha' > \alpha > 1$, it follows that $\lim_{x\to(-b/a)^+} Q_{s-1}(x) = -\infty$ and $\lim_{x\to+\infty} Q_{s-1}(x) = +\infty$, hence the sign variation of Q_{s-1} in $(-\frac{b}{a}, +\infty)$, which is equal to the sign variation of P'_{s-1}, is "$-, +$", thus the monotonicity of P_{s-1} is $\searrow\nearrow$. We have $\lim_{x\to+\infty} P_{s-1}(x) = +\infty$, so, if $P_{s-1}(-\frac{b}{a}) > 0$, the most sign varying possibility is "$+, -, +$", while if $P_{s-1}(-\frac{b}{a}) < 0$, the most sign varying possibility is "$-, +$". In either case, taking into account Lemma 26, the sign variation possibilities in $(-\frac{b}{a}, +\infty)$ for V_s are "$+, -, +$" or "$-, +$" or "$+$". As now, $V_s(-\frac{b}{a}) \le 0$, the actual sign variation for V_s is "$-, +$".

So, finally, the possible sign variations for V_s as x traverses from 0 to $+\infty$ are "$+, -, +$" or "$-, +$" or "$+$", so applying Theorem 22, the proof is concluded. ■

Proposition 33 *Let $\alpha' > 1 > \alpha > 0$ and $\theta_1, \theta_2 > 0$. The $W(\alpha', \theta_1)$ distribution is less $s-$IFR than the $W(\alpha, \theta_2)$ distribution.*

Proof. The argument is the same as that of the proof of Proposition 31. ■

5.3 About the role of iteration parameter

The criteria defined in Corollary 29 states a dependence on the iteration parameter s through the last term in both expressions. The applications discussed in the previous subsections have, actually, not used this dependence on s to characterize the order relation. This is due to the particular shape of the P_s or P_{s-1} functions constructed, showing relatively few sign variations and becoming infinite when $x \to +\infty$. We discuss now an example illustrating that this term may play a relevant role.

Example 34 *Let X have Pareto distribution with tail $\overline{F}_X(x) = \frac{1}{(x+1)^\alpha}$, and Y be the mixing of two Pareto distributions with tail $\overline{F}_Y(x) = \frac{p}{(x+1)^\alpha} + \frac{q}{(x+1)^{\alpha+1}}$, where $p, q > 0$ verify $p + q = 1$, and $\alpha > s - 1$, so the moments appearing in the definition of the $s-$iterated distributions (remember (4)) are finite. In order to use the criterium introduced in Corollary 29, denote $U(x) = \frac{\overline{F}_Y(x)}{\overline{F}_X(ax+b)}$, so we may rewrite*

$$P_{s-1}(x) = \log U(x) - (s-1)\log a + \log \frac{\mathbb{E}X^{s-1}}{\mathbb{E}Y^{s-1}}. \qquad (12)$$

It is easily seen that $\frac{\mathbb{E}X^{s-1}}{\mathbb{E}Y^{s-1}} = (1 - q\frac{s-1}{\alpha})^{-1}$ and $\lim_{x\to+\infty} U(x) = pa^\alpha$. Moreover, computing U', a little algebra shows that its sign is the same as the sign of $\ell(x) = (p(a-(b+1))\alpha + a(p-1))x + (a-(b+1))\alpha + (b+1)(p-1)$. To produce the convenient sign variations, we are interested in choosing the slope and the intercept of ℓ with opposite signs. We distinguish two different cases, depending on the sign of b.

Case 1: $b \geq 0$. *For this choice of the parameter* b, *we have that* $U(0) = (b+1)^\alpha \geq 1$, *hence* $\log U(0) \geq 0$. *Choose* $a > 1$ *such that* $pa^\alpha = (b+1)^\alpha$. *This means that* $a > b+1$ *and that*

$$P_{s-1}(0) = \lim_{x\to+\infty} P_{s-1}(x) = \alpha \log a + \log p - (s-1)\log a - \log\left(1 - q\frac{s-1}{\alpha}\right). \qquad (13)$$

It is obvious that P_{s-1} *is not constant and its monotonicity is the same as the monotonicity of* $U(x)$, *which is easier to determine. The only compatible choice is for the sign of the intercept and slope of* ℓ *is a positive intercept and a negative slope, although with different characterizing regions depending on the sign of* $p(\alpha+1)-1$. *Therefore, the sign of* U', *as* x *traverses from 0 to* ∞, *is "*$+,-$*", meaning that* $U(x)$ *is initially increasing and eventually decreasing. Hence, going back to* P_{s-1}, *this function will first be increasing and then decreasing, reaching at infinity the same value it has for* $x = 0$, *given the choice for* a. *Choose now* $a > 1$ *such that*

$$(\alpha - (s-1))\log a < \log \frac{\alpha - q(s-1)}{p\alpha}.$$

For such choice of a, *the sign variation of the function* P_{s-1} *is "*$-,+,-$*", which is compatible with* $Y \leq_{s-IFR} X$.

Case 2: $b < 0$. *Now we have that, for* $x < -\frac{b}{a}$, $U(x) = \overline{F}_Y(x) \leq 1$ *and* U *is decreasing, thus* P_{s-1} *is also decreasing for this choice of* x. *Moreover,* $P_{s-1}(0) = -(s-1)\log a - \log\left(1 - q\frac{s-1}{\alpha}\right)$. *For* $x \geq -\frac{b}{a}$, P_{s-1} *behaves as described in previous case. Therefore, for* $x \geq -\frac{b}{a}$, P_{s-1} *is first increasing and then decreasing. Taking into account that* $b < 0$, *we have that* $P_{s-1}(-\frac{b}{a}) \leq -(s-1)\log a + pa^\alpha$. *So, choosing conveniently the parameter* a, *we may obtain the sign variation, as* x *traverses from 0 to* $+\infty$, *for* P_{s-1} *as "*$+,-,+$*", which is compatible with* $X \leq_{s-IFR} Y$, *reversing the order relation obtained in the previous case.*

That is, this example shows that the iteration parameter s *may play an active role for the conclusion about the iterated order relation.*

Finally, remember Remark 19, for another example where the $s-$IFR comparison may depend on the iteration parameter s.

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